OpenGamma Quantitative Research

Forward CDS, Indices and Options

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Abstract

This paper is a followup to *The Pricing and Risk Management of Credit Default Swaps, with a Focus on the ISDA Model* [Whi13]. Here we show how to price indices (portfolios of CDSs) from the calibrated credit curves of the constituent names, and how to adjust those curves to match the market price of a index (basis adjustment). We then show how to price forward starting single-name CDSs and indices, since these are the underlying instruments for options on single-name CDSs and indices. The pricing of these options is the main focus of this paper. The model we implement for index options was first described by Pedersen [Ped03], and we give full implementation details and examples. We discuss the common risk factors (the Greeks) that are calculated for these options, given various ways that they may be calculated and show results for some example options. Finally we show some comparisons between our numbers and those displayed on Bloomberg’s CDSO screens. All the code used to generate the results in this paper is available as part of the open source release of the OpenGamma Platform.
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1 Introduction

In a previous paper [Whi13] we discussed how to price a single-name CDS, and gave particular attention to the ISDA model. In this paper we assume the reader is familiar with the mechanics of standard single-name CDS, and only clarify rather than repeat what is in other sources. We show how to build on a CDS pricing model\(^1\) to price forward CDS, CDS indices, options on single-name CDS (CDS options or default swaptions) and options on CDS indices (CDS index options).

In section 2 we briefly review standard single-name CDS and how they are priced from a known yield curve and credit curve. We do not discuss the calibration of credit curves from market quotes of CDS prices, as this is covered in our previous paper [Whi13] (in the context of the ISDA model). Section 3 introduces the forward starting CDS, and shows how (with minor modifications) these can be priced in the same way as a spot starting CDS. Section 4 discussed European options on single-name CDS (default swaptions), and shows that for certain contract specifications these can be priced within a Black framework. In section 5 we introduce portfolios of single-name CDS - \textit{CDS indices}, and show how to adjust individual credit curves so that the intrinsic price of the index matches the market. Section 6 then deals with European options on CDS indices, and in particular discusses Pedersen’s one factor model [Ped03] to price these options, as well as an approximation using the Black formula. Section 7 discusses numerical implementation and approximations and gives performance metrics, while section 8 introduces risk measures associated with index options - the Greeks. Finally section 9 discusses how options are priced from the Bloomberg CDSO screen, with examples comparing the results from those screens with our own calculations.

Appendix A gives a very brief overview of option pricing theory and the description of yield (discount) and credit curves, while appendix B gives a full list of terms used in this paper.

2 Review of Standard CDS

We use \textit{today} and \textit{trade date} interchangeably to mean the date on which we are valuing a trade,\(^2\) and assign this a value, \(t\). A spot CDS has some cost of entry (which can be positive or negative), which is paid on the \textit{cash settlement date},\(^3\) which we call \(t_{cs}\).

Standard CDS contracts have fixed coupons\(^4\) which are nominally paid every three months on the IMM dates (the 20\(^{th}\) of March, June, September and December). They are price quoted\(^5\) with a \textit{clean price}, and the actual cash payment (made on the cash settlement date) is this clean price (i.e. PUF × notional) adjusted for the \textit{accrued amount}. The accrued amount is the year fraction\(^6\) between the \textit{accrual start date} (the previous IMM date) and the \textit{step-in date} (one day after the trade-date) multiplied by the coupon; the convention is that this is quoted as a negative number for the buyer of protection and positive for the seller. Let \(\Delta\) be the year fraction and \(C\)

\(^1\)We do not restrict ourselves to the ISDA model.
\(^2\)Trade date is sometimes taken as the inception date of a derivative (a CDS or an option in this context), but we wish to value the derivative at any date up to expiry.
\(^3\)Usually three working days after the trade date.
\(^4\)These are 100 or 500 bps in North America and 25, 100, 500 or 100 bps in Europe.
\(^5\)Single name CDS are quoted with Points-Up-Front (PUF), which is the clean value as a percentage of notional; Indices are quoted as (100-PUF).
\(^6\)On a ACT/360 day count.
the coupon, then the cash settlement amount (for the buyer of protection) is
\[
\text{cash settlement amount} = \text{quoted price} + \text{accrued} = \text{quoted price} - C\Delta
\]
The buyer of protection pays this amount at \( t_{cs} \) to the seller of protection.\(^7\) The market value is the cash settlement amount, risk free discounted (the few days) from the cash settlement date back to today.

To clarify, clean price is the market quoted price, and dirty price is the cash settlement amount (clean price plus accrued), both of which are amounts ‘seen’ on the cash settlement date rather than the trade date.

### 2.1 Pricing from Yield and Credit Curves

We assume for the rest of this paper that there exists appropriate\(^8\) yield (discount) and credit curves. See appendix A for a brief overview of the curves and our previous paper [Whi13] for details of calibration (of the ISDA model). Where terms are common between this and the previous paper, we use the same notation - a full list of terms is given in appendix B.

#### 2.1.1 The Protection Leg

The value at time \( t_{cs}(\geq t) \) for protection between \( t \) and \( T \) where no default has occurred by \( t \) is
\[
V_{\text{Prot}}(t, T) = E_t\left[ e^{-\int_{t_{cs}}^{T} r_{\tau} d\tau} (1 - R(\tau)) I_{\tau \leq T} | \tau > t \right] = \frac{(1 - \hat{R}(T))}{P(t, t_{cs})} \int_{t}^{T} P(t, u) [-\hat{Q}(t, u)] du
\]
where \( \hat{R}(T) = E_t[R(\tau)] \) is the expected recovery rate (for a CDS with maturity \( T \)), \( \hat{Q}(t, u) \equiv \frac{dQ(t,u)}{du} \) and we have implicitly set the notional to unity. Where there is no term structure of recovery rates, we just use \( R \). The division by the discount factor \( P(t, t_{cs}) \) occurs because protection is from \( t \) but we value it at \( t_{cs} \).

#### 2.1.2 The Premium Leg

The full value (or dirty price) of the premium at \( t_{cs} \) is
\[
V_{\text{Prem}}(t, T) = \frac{C}{P(t, t_{cs})} \sum_{i=1}^{M} \Delta_i P(t, t_i) Q(t, e_i) + \eta_i \int_{\max(s_i, T_i)}^{e_i} (t - s_i) P(t, u) [-\hat{Q}(t, u)] du
\]
where \( C \) is the CDS coupon, \( s_i, e_i \) and \( t_i \) are the accrual start, end and payment time for the \( i \)th accrual period,\(^9\) and \( \eta_i \) is the day count adjustment factor.\(^10\) The sum is over all the accrual periods between \( t \) and \( T \), i.e. \( s_1 \leq t, t_1 > t \) and \( e_M = T \).

The clean price is then simply
\[
V_{\text{Prem}}(t, T) = V_{\text{Prem}}(t, T) + C\Delta
\]
where \( \Delta \) is the year fraction between \( s_1 \) and \( t + 1 \) (the step-in date).

\(^7\)Of course, if the amount is negative, the seller of protection actually pays the buyer.
\(^8\)The yield curve is for the currency in which the CDS is priced and the credit curve is for the reference entity.
\(^9\)Usually have \( e_i = t_i \), except for the last accrual period for which the accrual end is not business-day adjusted.
\(^10\)Ratio of the year fraction for the \( i \)th accrual period measured with the accrual day count convention to the same period measured with the curve day count convention. In almost all cases this is just 365/360.
From a mathematical sense, the dirty value (of the premium leg) is the principle quantity, since it is the actual economic value of future liabilities, and is computed directly from the yield and credit curves; the clean value is then found by adding the accrued amount.\(^{11}\) However the market treats the clean value as if it is the principle quantity, with the dirty value ‘calculated’ by subtracting the accrued amount. This also holds for other quantities that depend on the value of the premium leg, such as the price, annuity and spread, which are usually quoted as clean. This is perfectly sensible, since in reality the credit curve is not exogenously given, but calibrated from the market clean prices of CDS.

The \textit{risky annuity} or \textit{risky PV01} (RPV01) is the value of the premium leg per unit of spread, i.e. \(C = 1\) in the above equations.\(^{12}\) We will refer to it as simply the annuity. The relationship between clean and dirty annuity is clearly

\[ A(t, T) = AD(t, T) + \Delta \]  

where we have dropped to subscript for the clean annuity.

\subsection*{2.1.3 Price and Spread}

The clean price is given by

\[ V(t, T) = V_{\text{Prot}}(t, T) - V_{\text{Prem}}(t, T) = V_{\text{Prot}}(t, T) - CA(t, T) \]  

The spread is defined as the value of the coupon, \(C\) that makes the clean price zero. So

\[ S(t, T) = \frac{V_{\text{Prot}}(t, T)}{A(t, T)} \]  

This allows us to write the price as

\[ V(t, T) = [S(t, T) - C] A(t, T) \]  

where both the spread and the annuity are computed from the known yield and credit curves.

\subsection*{2.2 Constant Hazard Rate Pricing}

The convention for CDS post-'Big-Bang'\(^{13}\) is to quote the spread by pricing with a constant hazard rate; this has always been the convention for CDS indices. This spread is known as the \textit{quoted spread} or \textit{flat spread}, and differs from the par spread computed from the full credit curve (see White (2013) or Rozenberg (2009) [Whi13, Roz09] for further discussion).

If we set \(Q(t, T) = \exp(-(T-t)\lambda)\), the CDS price can be written explicitly as a function of the (constant) hazard rate, \(\lambda\),

\[ V(t, T|\lambda) = \tilde{V}_{\text{Prot}}(t, T|\lambda) - CA(t, T|\lambda) \]  

\(^{11}\)The accrued amount is \(C\Delta\) so is always positive (or zero).

\(^{12}\)Strictly the RPV01 is the value per basis point of spread, so is one ten-thousandth of the value per unit of spread.

\(^{13}\)In 2009 ISDA issued the ‘Big Bang’ protocol in an attempt to restart the market by standardising CDS contracts [Roz09].
where the ‘bar’ indicates quantities that have been calculated from a constant hazard rate. This is a monotonic function, so we may also find the inverse function\footnote{If the recovery rate, $R$ is too high, there may be no solution for a given $V$.}, given the clean price we can find $\lambda$ - this is the implied hazard rate for a given market price of a CDS. Similarly, we may write the spread as a function of $\lambda$

$$S(t, T|\lambda) \equiv \frac{V_{Prot}(t, T|\lambda)}{A(t, T|\lambda)}$$

or think of $\lambda$ as a function of $S$. This means we may write the annuity as

$$A(t, T|S) \equiv A(t, T|S(t, T)) = A(t, T|\lambda(S(t, T)))$$

which means: find $\lambda$ (by root finding) that gives this spread, then compute the annuity for that $\lambda$. The CDS price can now be expressed as an explicit function of spread

$$V(t, T|S) = (S - C) \times A(t, T|\lambda).$$

Figure 1 shows the annuity as a function of spread for a short (1Y) and long (10Y) expiry CDS. The annuity can be approximated extremely well by a low order polynomial.

![Risky Annuity vs. Spread](image.png)

Figure 1: The annuity as function of spread for a 1Y and 10Y CDS. All calculations are performed using a constant hazard rate.

It must be kept in mind that this a quoting convention only - $V(t, T) \leftrightarrow S(t, T)$ - the constant hazard rate ‘curve’ cannot be used to price any other maturity CDS. Also the annuity, protection
leg and spread calculated this way will generally not equal their equivalent values calculated from the full credit curve, i.e.

\[ V_{Prot}(t, T) \neq V_{Prot}(t, T) \]
\[ A(t, T|\bar{S}) \neq A(t, T) \]
\[ S(t, T) \neq S(t, T). \]

However, since the price must be the same whether a full or flat credit curve is used, it does follow that

\[ V_{Prot}(t, T) - CA(t, T|\bar{S}) = V_{Prot}(t, T) - CA(t, T) \]
\[ (S(t, T) - C)A(t, T|\bar{S}) = (S(t, T) - C)A(t, T). \] \hspace{1cm}(12)

3 **Forward Starting CDS**

A forward starting payer CDS entered into at time \( t \) will give protection against the default of an obligor for the period \( T_e > t \) to \( T_m \), in return for premium payments in that period. If the obligor defaults before the start of protection (i.e. \( \tau < T_e \)), the contract cancels worthless.

This can easily be replicated by entering a long protection CDS with maturity \( T_m \) and a short protection position with maturity \( T_e \) for the same obligor. Since for standard CDS, premiums are all paid on the same dates (adjusted IMM dates) with standard coupons, these will exactly cancel up to \( T_e \), leaving only the coupons between \( T_e \) and \( T_m \) to pay. Furthermore, if a default occurs before \( T_e \) the protection payments will exactly cancel.\(^{16}\) So the value of our forward CDS is given by

\[ V(t, T_e, T_m) = V(t, T_m) - V(t, T_e). \] \hspace{1cm}(13)

Since this value depends on the difference in price of a long and short dated CDS, its price will be sensitive to the shape of the credit curve.

In practice, CDS prices may not be available for a particular forward protection span,\(^{17}\) so we need to be able to price these instruments from the calibrated yield and (appropriate) credit curve.

### 3.1 Pricing from Curves

Assume that today, \( t \), I enter a forward contract where I agree that at some future expiration date, \( T_e \), I will enter a (payer) CDS contract with maturity \( T_m \) \((> T_e)\) in exchange for a payment of \( F - C\Delta \) made at the exercise settlement date \( t_{es} \geq T_e \) provided that the reference entity has not defaulted by \( T_e \). The exercise settlement date should correspond to the cash settlement date of a spot CDS entered at \( T_e \).\(^{18}\)

At \( T_e \), a spot starting CDS with maturity \( T_m \) will have a (clean) price for the cash settlement date \( t_{es} \) given by \( V(T_e, T_m) \). Therefore the mark-to-market value of the forward contract at \( T_e \) is

\[ P(T_e, t_{es}) [V(T_e, T_m) - F] \downarrow_{\tau>T_e} \]

\(^{15}\)This is not entirely true, since the final coupon period (for a standard CDS) has an extra day of accrued interest, so will not exactly match the coupon of the longer CDS.

\(^{16}\)This may need some legal specifications, so that the defaulted bond is delivered from the short position before it is delivered to the long position.

\(^{17}\)In particular if \( T_e \) does not correspond to an IMM date.

\(^{18}\)Normally three working days after \( T_e \).
and its expected value at \( t \) is
\[
Fwd(t, T_e, T_m) = \mathbb{E}_t \left[ \frac{\mathbb{I}_{t > T_e}}{\beta_t(T_e)} (V(T_e, T_m) - F) \right]
\]
where \( \beta_t(s) = \exp(\int_t^s r_u \, du) \) is the money market numeraire.\(^{19}\) The value of \( F \) that makes this zero is the forward (clean) price; it is given by
\[
F(t, T_e, T_m) = \frac{1}{P(t, T_e) Q(t, T_e)} \mathbb{E}_t \left[ \frac{\mathbb{I}_{t > T_e}}{\beta_t(T_e)} V(T_e, T_m) \right]
\]
\[
= \frac{1}{P(t, T_e) Q(t, T_e)} V(t, T_e, T_m)
\]
where \( V(t, T_e, T_m) \) is the expected present value of the forward CDS. This may be written as
\[
V(t, T_e, T_m) = \mathbb{E}_t \left[ \frac{\mathbb{I}_{t > T_e}}{\beta_t(T_e)} (V_{Prot}(T_e, T_m) - CA(T_e, T_m)) \right]
\]
\[
= V_{Prot}(t, T_e, T_m) - CA(t, T_e, T_m).
\]
We will show below how to calculate the expected present value of the protection leg and annuity from the calibrated yield and credit curves. In general we define
\[
Y(t, T_e) = \mathbb{E}_t \left[ \frac{\mathbb{I}_{t > T_e}}{\beta_t(T_e)} Y(T_e) \right]
\]
\[
= \mathbb{E}_t \left[ \frac{P(t, T_e) \mathbb{I}_{t > T_e}}{\beta_t(T_e)} Y(T_e) \right]
\]
(16)
to be the fair value today of a contract to pay \( Y(T_e) \) at \( t_{ea} \geq T_e \) providing the reference entity has not defaulted by \( T_e \). Implicit in the definition is the fact that the reference entity has not defaulted by \( t \). From the definition we have
\[
Y(T_e, T_e) = P(t, T_e) Y(T_e).
\]
(17)
A useful result is when \( Y(T_e) = 1 \), i.e. 1 unit of currency is paid at \( T_{ea} \) provided no default occurred before \( T_e \) - this a defaultable bond with a payment delay. We argue in appendix A that this is given by
\[
\mathbb{E}_t \left[ \frac{P(t, T_e) \mathbb{I}_{t > T_e}}{\beta_t(T_e)} \right] = P(t, T_e) Q(t, T_e).
\]
(18)
For a suitable choice of \( Y \), the quantity \( Y(t, T_e) \) is an asset which is positive provided that \( t < \tau \), so it can be used as a numeraire in what is known as a survival measure.\(^{19}\)

### 3.2 Valuing the Forward Protection

Protection between \( T_a \) and \( T_m \) can be statically replicated by buying protection to \( T_m \) and selling protection to \( T_e \). Therefore
\[
V_{Prot}(t, T_e, T_m) = P(t, T_e) (V_{Prot}(t, T_m) - V_{Prot}(t, T_e))
\]
\[
= (1 - \hat{R}(T_m)) \int_t^{T_m} P(t, u) [-\hat{Q}(t, u)] du
\]
\[
- (1 - \hat{R}(T_e)) \int_t^{T_e} P(t, u) [-\hat{Q}(t, u)] du.
\]
(19)
\(^{19}\)It is the amount of money in a ‘risk-free’ bank account at time \( s \geq t \) if one unit of currency in invested at \( t \).
If there is no term structure of expected recovery rates (i.e. \( \hat{R}(T_e) = \hat{R}(T_m) = R \)), then this simplifies to

\[
V_{\text{Prot}}(t, T_e, T_m) = (1 - R) \int_{T_e}^{T_m} P(t, u)[-\hat{Q}(t, u)] du, \tag{20}
\]

so the present value of the forward starting protection can be computed directly from the known yield and credit curve (which have been calibrated to spot instruments).

### 3.3 Valuing the Forward Premium Leg

The present value of the dirty forward starting annuity is

\[
A_D(t, T_e, T_m) \equiv \mathbb{E} \left[ e^{-\int_{t_0}^{t} r_s ds} \mathbb{1}_{t > T_e} A_D(T_e, T_m) \right] = \mathbb{E} \left[ e^{-\int_{t_0}^{t} r_s ds} \mathbb{1}_{t > T_e} (A(T_e, T_e) - \Delta) \right] = A(t, T_e, T_m) - P(t, t_{es})Q(t, T_e)\Delta. \tag{21}
\]

By writing this as \( A_D(t, T_e, T_m) = A_D(t, T_m) - A_D(t, T_e) \), or by evaluating the terms in the expectation directly, one can show that the present value of the forward starting annuity is given by

\[
A(t, T_e, T_m) = \sum_{i=1}^{M} \left[ \Delta_i P(t, t_i)Q(t, e_i) + \eta_i \int_{\max(s_i, T_e)}^{e_i} (t - s_i)P(t, u)[-\hat{Q}(t, u)] du \right] + P(t, t_{es})Q(t, T_e)\Delta \tag{22}
\]

where \( s_i, e_i \) and \( t_i \) are the accrual start and end times and payment time for the \( i \)th accrual period.\(^{20}\) The sum is over all the accrual periods between \( T_e \) and \( T_m \), i.e. \( s_1 \leq T_e, t_1 > T_e \) and \( e_M = T_m \).

### 3.4 The Forward Spread

The forward spread is the coupon that makes the expected value of the forward CDS zero, that is

\[
S(t, T_e, T_m) = \frac{V_{\text{Prot}}(t, T_e, T_m)}{A(t, T_e, T_m)}. \tag{23}
\]

Since the discounting from \( t_{es} \) to \( T_e \) cancels,

\[
S(T_e, T_e, T_m) = S(T_e, T_m) \tag{24}
\]

so the forward spread observed \( T_e \) is just the spot spread. The present value of a forward starting CDS can be written as

\[
V(t, T_e, T_m) = \mathbb{E}_t \left[ \frac{\mathbb{1}_{t > T_e}}{\beta(t_{es})} (S(T_e, T_m) - C)A(T_e, T_m) \right] = A(t, T_e, T_m)\mathbb{E}_t^{\hat{A}} [S(T_e, T_m) - C]. \tag{25}
\]

\(^{20}\)As we mentioned earlier, usually have \( e_i = t_i \), except for the last accrual period for which the accrual end is not business-day adjusted.
In the last line we have switched to the annuity measure (see appendix A for details). Since this can also be written as

\[ V(t, T_e, T_m) = [S(t, T_e, T_m) - C] A(t, T_e, T_m) \]  

(26)

we have that \( E_t^A [S(T_e, T_m)] = S(t, T_e, T_m) \), that is, the forward spread is a Martingale in the annuity measure.\(^{21}\)

### 3.4.1 Different Expressions for Forward Spread

The forward spread can also be expressed as

\[ S(t, T_e, T_m) = \frac{V_{Prot}(t, T_m) - V_{Prot}(t, T_e)}{A(t, T_m) - A(t, T_e)} \]

\[ = \frac{S(t, T_m)A(t, T_m) - S(t, T_e)A(t, T_e)}{A(t, T_m) - A(t, T_e)} \]

\[ = S(t, T_m) + \frac{A(t, T_e)}{A(t, T_m) - A(t, T_e)} (S(t, T_m) - S(t, T_e)) \]  

(27)

The final expression shows that, without explicitly calculating the annuities, if the term structure of spreads is upwards sloping (i.e. \( S(t, T_m) > S(t, T_e) \)), then the forward spread is greater than the spot spread (both with maturity \( T_m \)); the converse is true for a downwards sloping term structure of spreads.\(^{22}\)

### 3.5 Forward Flat Spread

In the previous section we showed how to calculate a forward par spread, which depends on the full credit curve. The concept of a forward flat spread is less well defined.

We may write the present value of a forward starting CDS in terms of a flat spread (and annuity) as

\[ V(t, T_e, T_m) = E_t \left[ \frac{P(T_e, t_{ex})I_{T_e > T_m}}{\beta(t_e)} (S(T_e, T_m) - C) \ A(T_e, T_m|\bar{S}) \right]. \]  

(28)

The present value of the (forward) flat annuity is

\[ A(t, T_e, T_m) = E_t \left[ \frac{P(T_e, t_{ex})I_{T_e > T_m}}{\beta(t_e)} \ A(T_e, T_m|\bar{S}) \right], \]  

(29)

however, calculation of this value depends on the distribution of \( \bar{S} \), so it cannot be calculated from the calibrated credit curve at \( t \). Additionally, the flat annuity is not an asset in its own right and we cannot technically use it as a numeraire (even if we could calculate its value). We may however switch to the risky bond measure to write

\[ F(t, T_e, T_m) = E_t^B \left[ (S(T_e, T_m) - C) \ A(T_e, T_m|\bar{S}) \right] \]  

(30)

\(^{21}\)This also follows from the fact that the spread is the ratio of an asset (the protection leg) to the numeraire (the annuity).

\(^{22}\)Regardless of the shape of the credit curve, we always have \( A(t, T_m) > A(t, T_e) \iff T_m > T_e \).
where the forward is defined in equation 14. Whatever the terminal distribution of \( \bar{S} \) is in the risky bond measure, \( \mathbb{B} \), it must satisfy the above equation. The standard approach is to assume it is log-normal, so

\[
\bar{S}(T_e, T_m) = S_0 \exp(Z \sigma \sqrt{T_e - t} - \sigma^2 (T_e - t)/2)
\]

(31)

where \( \sigma \) is the spread volatility, \( Z \) is a standard Gaussian random variable and \( S_0 \) is a constant we must compute. The value of \( S_0 \) is such that

\[
F(t, T_e, T_m) = \int_{-\infty}^{\infty} (\bar{S} - C) \bar{A}(T_e, T_m|\bar{S}) e^{-z^2/2} dz
\]

where \( \bar{S} = S_0 \exp(z \sigma \sqrt{T_e - t} - \sigma^2 (T_e - t)/2) \)

(32)

and \( \bar{A}(T_e, T_m|\bar{S}) \) is calculated at \( t \) assuming deterministic interest rates. The integral can be commuted via Gauss-Hermite quadrature \([PTVF07]\), and the value of \( S_0 \) found by one-dimensional root finding (again \([PTVF07]\)). Clearly

\[
\bar{S}_0 = \mathbb{E}^B [\bar{S}(T_e, T_m)]
\]

(33)

so \( \bar{S}_0 \) is the expected value (under \( \mathbb{B} \)) of the forward flat spread, but it does depend on the extraneous parameter \( \sigma \) (spread volatility).

### 3.5.1 Flat ATM Forward Spread

A useful quantity for the options we will meet later, is the value of \( \bar{S} \) that solves

\[
F(t, T_e, T_m) = (\bar{S} - C) \bar{A}(T_e, T_m|\bar{S}).
\]

(34)

This is of course just \( S_0 \) for \( \sigma = 0 \). We call this the flat ATM forward spread, \( K_{ATM} \). This is simply a number that put in a formula\(^{23}\) gives the correct forward price. It is not the expected value of the forward flat spread.

### 4 Default Swaptions

A default swaption is a European option that can be exercised at time \( T_e > t \) to enter into a CDS for protection between \( T_e \) and \( T_m \). A payer option exercises into a payer CDS and of course a receiver option exercises into a receiver CDS.

In a knockout option, if the reference entity defaults before the option expiry, the option is cancelled and there are no further cash flows. For a non-knockout payer default swaption, in the case of a default before expiry, the option holder would receive the default settlement amount.\(^{24}\) It should be noted that since the crisis, options on single-name CDS are not a widely traded instrument, and the material below really serves as a precursor to the discussion of options on CDS indices (which are traded).

\(^{23}\)The ISDA quoted spread to upfront calculation.

\(^{24}\)In practice they would deliver a suitable defaulted bond at expiry in exchange for par. In some circumstances this would be cash settled instead.
4.1 Spread based Default Swaptions

In this case, if an option strike at a level $K$ is exercised at $T_e$, one entered a bespoke CDS paying a coupon of $K$ with accrual starting from the day after the exercise date (i.e. $T+1$ so $\Delta = 0$), that is, the underlying is a pre-‘Big-Bang’ style CDS. When there was a market for options on single-name CDS, this was the standard contract specification.

The payoff in this case is

$$V_{\text{option}}(T_e, T_m) = P(T_e, t_{es})I_{T > T_e}A(T_e, T_m)(\chi [S(T_e, T_m) - K])^+$$

where $\chi = 1$ for a payer and $\chi = -1$ for a receiver option. The value of the option at time $t$ is given by

$$V_{\text{option}}^{\text{knockout}}(t, T_e, T_m) = \mathbb{E}_t \left[ \frac{P(T_e, t_{es})A(T_e, T_m)I_{T > T_e}(\chi [S(T_e, T_m) - K])^+}{\beta_t(T_e)} \right].$$

We may then use the annuity, $A(t, T_e, T_m)$, as a numeraire (see appendix A) and perform a change of measure, to write the option value as

$$V_{\text{option}}^{\text{knockout}}(t, T_e, T_m) = A(t, T_e, T_m)\mathbb{E}^A_t \left[ (\chi [S(T_e, T_m) - K])^+ \right]$$

where $A$ indicates we are in the risky annuity measure. As we have already discussed, the forward spread is a martingale in this measure.

The above argument is mathematically identical to that made for pricing interest rate swaptions. Any model that preserves the martingale property of the forward spread will produce arbitrage free option prices. The usual approach is to assume log-normal dynamics for $S$, i.e.

$$\frac{dS(u, T_e, T_m)}{S(u, T_e, T_m)} = \sigma(u) dW_u$$

which gives

$$V_{\text{option}}^{\text{knockout}}(t, T_e, T_m) = A(t, T_e, T_m)\text{Black}(S(t, T_e, T_m), K, T_e - t, \hat{\sigma}, \chi)$$

where

$$\hat{\sigma} = \sqrt{\frac{1}{T_e - t} \int_t^{T_e} \sigma(u)^2 du}$$

$$\text{Black}(F, K, T, \sigma, \chi) = \chi (F\Phi(d_1) - K\Phi(d_2))$$

$$d_1 = \ln \left( \frac{F}{K} \right) + \frac{\sigma^2 T}{2} \quad d_2 = d_1 - \sigma \sqrt{T}$$

is Black’s formula [Bla76]. The ATM strike is given by $K_{ATM} = S(t, T_e, T_m)$.

While it is attractive to end up pricing with the familiar Black formula, there is no market for the underlying instrument, which makes hedging non-trivial, and therefore it is unlikely that this style of default swaption will trade in the future.
4.2 Price Based Default Swaptions

In order to have a liquid instrument as the underlying, the option must be price-based, which means, at \( T_e \) the option holder chooses whether to excise the option and pay the excise price, \( G(K) \), in return for entering a standard CDS. The excise decision is made at \( T_e \) but the excise price is paid on the excise settlement date, \( t_{es} \) which corresponds to the settlement date of the underlying CDS. The excise price, \( G(K) \), is clean, but what is actually paid at \( t_{es} \) is the dirty price, \( G(K) - C\Delta \). This means the payoff at \( T_e \) is

\[
V_{\text{payer knockout}}(T_e) = P(T_e, t_{es}) \left[ (S(T_e, T_m) - C)A(T_e, T_m) - C\Delta) - (G(K) - C\Delta) \right]^+ \quad (41)
\]

The accrual terms have cancelled in the payoff, but one must remember that the actual amount paid to exercise the option does include the accrued - this seems to be overlooked in the literature. The choice of \( G(K) \) affects how simple this type of option is to price.

4.2.1 The Simple Case

The simplest case is when \( G(K) = (K - C)A(T_e, T_m) \). That is, the payment exactly accounts for the difference in value of the premium leg for paying the actual coupon, \( C \), rather than the nominal coupon, \( K \). This gives a payoff that is identical to equation 35, and thus is priced in the same way. However, it is highly unlikely to have a exercise price of this form since it is not known until expiry, and depends on a subjective calculation of the annuity at \( T_e \). Furthermore, as the exercise price is itself stochastic, one cannot simply hedge with the standard CDS of the same maturity as the underlying.

4.2.2 The Fixed Excise Price Case

The convention for index options is to set

\[
G(K) = (K - C)\overline{A}(T_e, T_m|K)
\]

that is, the annuity is computed using a spread level of \( K \), which gives a payoff that is zero for \( S(T_e, T_m) \leq K \) and is positive for \( S(T_e, T_m) > K \), so \( K \) is still a strike level. Technically \( \overline{A}(T_e, T_m|K) \) depends on the yield curve observed at \( T_e \), but in practice it is calculated on the trade date and treated as a known fixed payment. For some options (indices that trade on price) the excise price is given directly rather than quoted as a strike (which is then converted to a price). For the general case of an excise price \( G \) that is know at \( t \), the option price is

\[
V_{\text{option knockout}}(t, T_e, T_m) = \mathbb{E}_t \left[ \frac{P(T_e, t_{es})}{\beta_t(T_e)} \left( (S(T_e, T_m) - C) \overline{A}(T_e, T_m|S) - G \right)^+ \right] \quad (42)
\]

where we have expressed the CDS price in terms of flat spread and annuity, and switched to the risky bond measure (see appendix A). Note, the survival probability, \( Q(t, T_e) \), in the above expression is obtained from the full credit curve.
To compute the option price it is necessary to know the terminal distribution of $\tilde{S}$ in this measure. If we assume log-normal dynamics as we did in section 3.5, then once $S_0$ is calibrated the option price is easily calculated as

$$V_{\text{payer knockout}}(t, T_e, T_m) = P(t, t_{es})Q(t, T_e) \int_{-\infty}^{\infty} (\tilde{S} - C) \tilde{A}(\tilde{S}) - G)^+ e^{-z^2/2} dz$$

$$= P(t, t_{es})Q(t, T_e) \int_{z^*}^{\infty} (\tilde{S} - C) \tilde{A}(\tilde{S}) - G) e^{-z^2/2} dz$$

with $\tilde{S} = S_0 \exp(\sigma \sqrt{T_e - t} - \sigma^2(T_e - t)/2)$

where $z^*$ solves $(\tilde{S} - C) \tilde{A}(\tilde{S}) = G$. This integral must also be computed numerically, but since the payoff is not smooth, it is unwise to use Gauss-Hermite quadrature. It is better to use the second form with Simpson’s rule or a higher order scheme. One further point to note is that while function $\tilde{A}(\tilde{S})$, which appears in the integrand, is smooth, it is computed via a root finding routine, therefore in implementations it is often approximated. This is discussed further in section 7.

### 4.2.3 Put-call Parity

The value of payer minus a receiver option is given by

$$V_{\text{payer knockout}}(t, T_e, T_m) - V_{\text{receiver knockout}}(t, T_e, T_m) = \mathbb{E}_t \left[ \frac{P(T_e, t_{es})I_{t_{es} > T_e}}{\beta(T_e)} (V(T_e, T_m) - G(K)) \right]$$

$$= P(t, t_{es})Q(t, T_e) (F(t, T_e, T_m) - G(K)),$$

so the ATM exercise price is given by $G(K_{ATM}) = F(t, T_e, T_m)$, and the flat spread that solves this, is what we called the flat ATM forward spread.

### 4.3 No-knockout Options

A payer option can be structured without the knockout feature: In the event of a default before expiry, the option holder may deliver an appropriate defaulted bond at the expiry date in return for par. The value of this front-end protection is

$$V_{\text{FEP}}(t, T_e) = (1 - R)(1 - Q(t, T_e))P(t, t_{es})$$

where the default settlement occurs at $t_{es}$. If the default is settled immediately rather than waiting to the option expiry, the front end protection is just worth the same as the protection leg of a (spot) CDS with maturity of $T_e$.\footnote{In a low interest rate environment and with expiries typically less than six months, the difference is tiny.}

If there is a default before expiry, then the front-end protection pays out and the rest of the option is worthless. If there is no default before expiry, the option is worth the same as the no knockout. The front-end protection is worthless for the holder of a receiver option (who would never choose to pay out in the event of a default). So we have

$$V_{\text{payer no-knockout}}(t, T_e, T_m) = V_{\text{FEP}}(t, T_e) + V_{\text{payer knockout}}(t, T_e, T_m).$$
A complication to this picture is if the contract specifies that the excise price $G$ is paid in exchange for default settlement. In this case the payoff is\footnote{The accrual still cancels provided accrual is paid on default.}

\[
(\chi(V(T_e, T_m)\mathbb{1}_{\tau > T_e}) + (1 - R(\tau))\mathbb{1}_{\tau \leq T_e} - G))^+.
\]

It is not possible to decompose this into a knock-out option plus front-end protection. Also, the no-knockout receiver is worth more than the knockout receiver since the option holder will excise following a default provided that $G > (1 - R(\tau))$, where $R(\tau)$ is the realised recovery rate. The option price can be decomposed as

\[
V_{\text{no-knockout}}(t, T_e, T_m) = V_{\text{knockout}}(t, T_e, T_m) + \mathbb{E}_t \left[ \frac{P(T_e, t_{es})\mathbb{1}_{\tau \leq T_e}}{\beta_t(T_e)} (\chi(1 - R(\tau) - G))^+ \right]. \tag{47}
\]

Computation of the second term requires a model for the realised recovery rate. A treatment is given in Martin(2012) [Mar12].

\section{CDS Indices}

CDS Indices are portfolios of single-name CDS, with typically between 40 and 125 entries. They pay coupons every three months (i.e. on adjusted IMM dates) to the buyer of the index (who is the seller of protection), and in return, the index buyer must receive any (qualifying) defaulted bonds in return for par. For constancy with single-name CDS, we use \textit{payer} to mean the payer of coupons (i.e. the seller of the index/buyer of protection), and \textit{receiver} to mean the recipient of coupon payments (i.e. the buyer of the index/seller of protection).

The names in the index are given a weight, with $w_j$ the weight of the $j$th entry. If the notional of a index is $N$, then the notional of the $j$th entry is just $N_j = N w_j$, and

\[
\sum_{j=1}^{J} w_j = 1 \tag{48}
\]

where $J$ is number of names in the index. Most indexes are equally weighted, so $w = \frac{1}{J}$.

Let $J_D(t)$ be the number of names that have expired by some time $t$, then

\[
\sum_{j=1}^{J} \mathbb{1}_{\tau_j > t} w_j = f(t) < 1 \iff J_D(t) > 0 \tag{49}
\]

where $f(t)$ is the index factor (at time $t$), and $\tau_j$ is the default time of the $j$th name; so the index factor is less than one iff any names in the index have defaulted.

The amount paid on the $i$th coupon payment date is

\[\text{NC}_i f(t_i)\]

where $C_i$ is the index coupon and $\Delta_i$ is the year fraction for the $i$th accrual period.

If the $j$th name defaults at $\tau_j$, the nominal amount paid out at $\tau_j$ is

\[N w_j (1 - R_j)\]
where $R_j$ is the realised recovery rate for that name. If this default occurs in the $i^{th}$ accrual period, the buyer of protection must pay the accrual-on-default of

$$Nw_j \Delta(s_i, \tau_j)C$$

where $\Delta(s_i, \tau_j)$ is the year fraction between the accrual start date and the default. So the cash settlement for the default is

$$Nw_j ((1 - R_j) - \Delta(s_i, \tau_j)C).$$

Provided that the definition of a default is the same, the payments above will exactly match those from owning the CDS (for the same maturity) on all the individual constituent of the index, each with notionals of $N_j = Nw_j$. So the index can, in principle, be statically replicated.

### 5.1 Index Value

The price of an index is given as a percentage in a form such that the cash settlement amount at $t_{cs} \geq T$ is

$$\text{cash settlement amount} = N \times f \times (1 - \text{quoted price} - C \Delta).$$

So the index quoted price and the points upfront (PUF) are related by

$$1 - \text{quoted price} = \text{PUF}$$

For the remainder of this paper we will assume that the index quoted price is expressed as PUF.

CDS indices are conventionally priced using a constant hazard rate, $\lambda$. This means the index annuity at time $t$ for an index that expires at $T$ may be written as $A_I(t, T|\lambda)$ and computed exactly as in the single-name case. The value of the protection leg (for a unit notional) is given by

$$V_{I; \text{Prot}}(t, T|\lambda) = (1 - R_I) \int_t^T P(t, s) e^{-\lambda(s-t)} ds \quad (50)$$

where $R_I$ is the index recovery rate, which is a fixed number with little relation to the recovery rates of the constituent names.\(^{29}\) The index PUF is

$$PUF_I(t, T) = V_{I; \text{Prot}}(t, T|\lambda) - CA_I(t, T|\lambda) \quad (51)$$

and its value (for a payer) is

$$V_{I; \text{payer}}(t, T) = Nf(t) PUF_I(t, T) \quad (52)$$

where $f(t)$ is the index factor at time $t$, and the protection and annuity are computed for unit notional. The index value will change sign depending on whether one is considering the payer or receiver, but the PUF is always given as shown above.

For a given index PUF, we can find the implied hazard rate $\lambda$. The maximum value of the protection leg is $(1 - R_I)$ (i.e. immediate default) which is also the maximum value of the index PUF, so if the PUF of the index is above $(1 - R_I)$ it is not possible to find a hazard rate and thus an equivalent spread.\(^{30}\)

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\(^{27}\)which is also the previous payment, except for the first accrual period.

\(^{28}\)The standard is three working days.

\(^{29}\)It is usually 40%, but is 30% for CDX.NA.HY and other indices made of lower credit quality entities. Basically it is a number that lets you convert between PUF and a quoted spread.

\(^{30}\)This is just an effect of using a recovery rate that is too high.
5.2 Index Spread

The index spread is defined as

\[
S_I(t; T_j) = \frac{V_{I,Prot}(t; T_j)}{A_I(t; T_j)}
\]

where we have made it explicit that it is a function of the hazard rate (which in turn is implied from the index price). For a given \(\lambda\) this is independent of the index factor, \(f(t)\).

Given a spread, one can also imply the hazard rate that satisfies the above equation. So, like the single-name case, the annuity can be expressed as a function of the spread, \(\hat{A}_I(t; T_j|\hat{S}_I)\), and the index PUF and value may be written as

\[
PUF_I(t; T_j|\hat{S}_I) = (\hat{S}_I(t; T_j) - C)\hat{A}_I(t; T_j|\hat{S}_I)
\]

\[
V_{I,\text{payer}}(t; T_j|\hat{S}_I) = N f(t)(\hat{S}_I(t; T_j) - C)\hat{A}_I(t; T_j|\hat{S}_I).
\]

It is worth mentioning again that the (flat) annuity function \(\hat{A}_I(t; T_j|\hat{S}_I)\) as we have defined it, is always for a unit notional index: This is so it can be regarded as a function of the index (flat) spread, \(\hat{S}_I\), only, and not other state variables.\(^{31}\)

An index value may be quoted as a PUF or a (flat) spread. Just as in the single-name CDS case, the conversion is via a constant hazard rate (i.e. the ISDA upfront model).

5.3 Intrinsic Value

Since an index is a portfolio of single-names, we can value it using the individual credit curves of the constituent names; this is the intrinsic value. At some time, \(t\), the clean value of the index (for the protection buyer/seller of the index) is

\[
V_{I,\text{intrinsic}}(t; T) = N \sum_{j=1}^{J} I_{I_j > t} w_j \left[ V_{Pr, j}(t; T) - C A_j(t; T) \right]
\]

\[
= N \sum_{j=1}^{J} I_{I_j > t} w_j (S_j(t; T) - C) A_j(t; T)
\]

\[
= N \sum_{j=1}^{J} I_{I_j > t} w_j (\hat{S}_j(t; T) - C) \hat{A}_j(t; T|\hat{S}_j).
\]

The last form shows that we can compute this quantity using the flat spreads of the constituent names. Of course spreads (of either form) are not defined for defaulted names, but the default indicator means they are not counted, and in practice the sum is taken over just the undefaulted names.

It is unlikely that this value will equal the market quoted value - this is known as the index basis. The biggest effect will come from a difference between what constitutes a credit event in the index and for a single-name CDS: In North America restructuring is not a credit event for indices while it is for standard CDS [O’k08]. The second effect is liquidity: An index will be more liquid than its constituent parts, so it is more expensive to buy protection on all the individual names than to sell the index (for the same total notional) [Ped03].

\(^{31}\)Obviously it also depends on the payment schedule of the premium leg, but that is fixed.
5.4 Intrinsic Annuity and Spread

The intrinsic value of the index (eqn. 55) may be written in the form

\[ V_{I;\text{intrinsic}}(t, T) = N \left( \frac{\sum_{j=1}^{J} I_{\tau_j > t} w_j S_j(t, T) A_j(t, T)}{\sum_{j=1}^{J} I_{\tau_j > t} w_j A_j(t, T)} - C \right) \sum_{j=1}^{J} I_{\tau_j > t} w_j A_j(t, T) \]

Comparing this with equation 54, we see that we can define the intrinsic annuity and intrinsic spread as

\[ A_{I;\text{intrinsic}}(t, T) = \sum_{j=1}^{J} I_{\tau_j > t} w_j A_j(t, T) \]  
\[ S_{I;\text{intrinsic}}(t, T) = \frac{\sum_{j=1}^{J} I_{\tau_j > t} w_j S_j(t, T) A_j(t, T)}{\sum_{j=1}^{J} I_{\tau_j > t} w_j A_j(t, T)} \]

The difference between the index spread and the intrinsic spread is another way of seeing the index basis. These definitions give the intrinsic index value as

\[ V_{I;\text{intrinsic}}(t, T) = N(S_{I;\text{intrinsic}}(t, T) - C)A_{I;\text{intrinsic}}(t, T). \]  

The form is identical to equation 54 except there is no explicit index factor - this is because it is implicit in the definition of the intrinsic annuity, \( A_{I;\text{intrinsic}}(t, T) \). These intrinsic values could alternatively be calculated using flat spreads and annuities, which would result in slightly different values. However, since we are ultimately interested in computing forward index values, we will always work with values computed from the full credit curves of the constituent names.

5.5 Credit Curve Adjustment

The individual credit curves will be used to calculate a forward price for the index, and in turn price options on the index. As with all models to price derivatives, it is essential that the model prices the underlaying exactly. To this end we need to adjust the individual credit curves in some way so that the intrinsic value matches the market value. There are several ways to do this (methods are discussed in Pedersen [Ped03] and O’Kane [O’k08]).

5.5.1 Matching a Single Index

Since there is only one number, the market price, to match, we can only have a single adjustment factor. We choose to adjust the forward hazard rates directly, and apply a constant multiplication factor to the rate. Each credit curve is transformed thus:

\[ \tilde{h}_j(t, T) = \alpha h_j(t, T) \]
\[ \tilde{Q}_j(t, T) = (Q_j(t, T))^\alpha \]
\[ \tilde{\Lambda}_j(t, T) = \alpha \Lambda_j(t, T) \]

where \( \alpha = \mathcal{O}(1) \). Since the ISDA standard model uses log-linear interpolation of survival probabilities, these curve properties are preserved by applying the multiplication factor to the value of the knots, i.e. \( \tilde{\Lambda}_j^i = \alpha \Lambda_j^i \), where \( \Lambda_j^i \) is the \( i \)th knot of the \( j \)th curve.

\[^{32}\text{we could normalise the intrinsic annuity (i.e. divide by } f(t) \text{) so that the index factor is explicit in the equal for an intrinsic value, but we find this more cumbersome.}\]
\[^{33}\text{A multiplicative factor is preferred over an additive one, since it ensures that forward hazard rates remain positive, and is more in keeping with the dynamics of hazard rates - volatility tends to scale with the rate.}\]
\[^{34}\text{Equivalently, piecewise constant hazard rates.}\]
We may write the adjusted intrinsic value as a function of $\alpha$, $V_{I,\text{intrinsic}}(t, T|\alpha)$, and find value of $\alpha$ that satisfies

$$V_{I,\text{intrinsic}}(t, T|\alpha) = V_{I,\text{market}}(t, T)$$

using a one-dimensional root finder.

We could, instead, make the adjustment so that the intrinsic spread matches the index spread. However the index values would not (exactly) match, and since the spread is just a quoting convention - it is the index value that matters - this is not desirable.

### 5.5.2 Matching Multiple Indices

What we mean here, is matching several indices that have the same constituents but different expiries (e.g. iTraxx Europe 3Y, 5Y, 7Y and 10Y). In general we could transform the forward hazard rates of each credit curve thus:

$$\tilde{h}_j(t, T) = \alpha(T)h_j(t, T)$$

$$\tilde{Q}_j(t, T) = \exp\left(-\int_t^T \alpha(u)h_j(t, u)du\right)$$

$$\tilde{\Lambda}_j(t, T) = \frac{1}{T-t} \int_t^T \alpha(u)h_j(t, u)du$$

and find the single curve $\alpha(T)$ such that the intrinsic price for each index equals the market value. The simplest way to do this is to make $\alpha(u)$ piecewise constant, i.e.

$$\alpha(u) = \alpha_i \text{ for } T_{i-1} \leq u < T_i$$

where $T_i$ is the maturity of the $i^{th}$ index. We can then bootstrap $\alpha$ by solving for each index maturity in turn.

For the ISDA model this turns out to be quite easy. The first step is to add additional knots to the credit curves at the index expiries (if these knots are not already present), at a level interpolated from the existing points - this has no effect on any survival probabilities computed off the curves.

The next step is to multiply all the knots less than or equal to $T_1$ on all the curves by $\alpha_1$, then all the knots greater that $T_{i-1}$ and less than or equal to $T_i$ by $\alpha_i$ and finally all the knots greater than $T_{n-1}$ by $\alpha_n$, where $n$ is the number of index maturities. The values of $\alpha$ are found in turn by solving for the index value.

We finish this section with an example of this adjustment in practice. We choose the CDX.NA.HY.21-V1 index on 13-Feb-2014. The rates used to build the yield curve and the index prices are given in table 10. Data on par spreads at standard pillars and recovery rates for the 100 constituents are provided by Markit - we do not reproduce them here. With this data we can build the 100 individual credit curves and calculate the intrinsic price of the index at four different terms. We show the market and intrinsic prices in table 1 together with the equivalent quoted spreads. For the most liquid 5Y index the difference is less than one basis point between the market and intrinsic quoted spread.

Once we adjust the curves as described above, of course the adjusted intrinsic prices will match the market prices. It is instructive to see what effect this has on the par spreads of the

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35 An $x$ year CDS has maturity $x$ years from the next IMM date, while a $x$ year index has maturity $x$ years plus three months from the previous roll date, so the two maturities may not match.
Table 1: The intrinsic and market prices for the CDX.NA.HY.21-V1 index on 13-Feb-2014.

<table>
<thead>
<tr>
<th>Term</th>
<th>PUF Intrinsic</th>
<th>Quoted Spread Intrinsic</th>
<th>Market</th>
</tr>
</thead>
<tbody>
<tr>
<td>3Y</td>
<td>-7.248</td>
<td>234.770</td>
<td>223.955</td>
</tr>
<tr>
<td>5Y</td>
<td>-7.591</td>
<td>323.427</td>
<td>322.789</td>
</tr>
<tr>
<td>7Y</td>
<td>-6.576</td>
<td>381.069</td>
<td>396.038</td>
</tr>
<tr>
<td>10Y</td>
<td>-7.433</td>
<td>394.141</td>
<td>406.448</td>
</tr>
</tbody>
</table>

In table 2 below we show original and adjusted par spreads\(^{36}\) for four names in the index with very different credit qualities and spread term structures.

Table 2: The original par spreads, and the implied par spreads after credit curve adjustment for three names in the CDX.NA.HY.21-V1 index. Note, we only show three pillar points (6M, 5Y and 10Y) for space.

<table>
<thead>
<tr>
<th></th>
<th>Original Par Spreads</th>
<th>Adjusted Par Spreads</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>6M</td>
<td>5Y</td>
</tr>
<tr>
<td>The AES Corp</td>
<td>19.0</td>
<td>187.0</td>
</tr>
<tr>
<td>Bombardier Inc.</td>
<td>29.0</td>
<td>319.0</td>
</tr>
<tr>
<td>J C Penney Co Inc</td>
<td>1594</td>
<td>1633</td>
</tr>
<tr>
<td>TX Competitive Elec Hldgs Co LLC</td>
<td>47657</td>
<td>31568</td>
</tr>
</tbody>
</table>

In this example, the spreads at the long end (5 years and more) increase while the spreads at the short end decrease - this would be expected from what we saw in table 1. The exception to this is the highly distressed TX Competitive Elec Hldgs Co, which has a one year survival probability of just 3.6%; in this case the 10Y spread is actually driven by the short end of the credit curve, which is adjusted down.

For the remainder of this note, unless explicitly stated, it is assumed that the intrinsic credit curves have been adjusted to match index prices before any values are computed from them. With that in mind we will refer to the intrinsic index spread and annuity as just \(S_I(t,T)\) and \(A_I(t,T)\) - these should not be confused with the flat index spread and annuity, \(\tilde{S}_I(t,T)\) and \(\tilde{A}_I(t,T|S)\).

### 6 Options on CDS Indices

Without loss of generality, we set the (initial) index notional to one (i.e. \(N = 1\)) and drop the term.

\(^{36}\)Recall we adjust the credit curves directly, so these are the par spreads implied from the new credit curves.
6.1 Forward CDS Index

The (clean) value of the index observed at some time $T_e > t$ and valued at $t_{es}$ is

$$V_f(T_e, T_m) = f(T_e)(S_f(T_e, T_m) - C)\tilde{A}_f(T_e, T_m|S_f)$$

$$= (S_f(T_e, T_m) - C)A_f(T_e, T_m).$$  \hfill (61)

We now consider the present value of this index using the (adjusted) credit curves of the constituent names. This gives

$$V_I(t; T_e; T_m) = E_t \left[ \frac{P(T_e, t_{es})}{\beta(T_e)} \sum_{j=1}^{J} 1_{\tau_j > T_e} w_j A_j(t, T_e, T_m) (S_j(T_e, T_m) - C) \right]$$

$$= \sum_{j=1}^{J} \left( w_j E_t \left[ 1_{\tau_j > T_e} P(T_e, t_{es}) A_j(T_e, T_m) (S_j(T_e, T_m) - C) / \beta(T_e) \right] \right)$$

$$= \sum_{j=1}^{J} w_j A_j(t, T_e, T_m)(S_j(t, T_e, T_m) - C)$$

$$= \sum_{j=1}^{J} w_j V_{prot,j}(t, T_e, T_m) - C \sum_{j=1}^{J} w_j A_j(t, T_e, T_m).$$ \hfill (62)

Either of the last two forms are readily calculated from the (adjusted) calibrated credit curves of the undefaulted constituents at $t$.

6.2 Forward Spread

We may define the forward intrinsic protection leg and annuity as

$$V_{prot,f}(t, T_e, T_m) = \sum_{j=1}^{J} w_j V_{prot,j}(t, T_e, T_m)$$

$$A_f(t, T_e, T_m) = \sum_{j=1}^{J} w_j A_j(t, T_e, T_m)$$ \hfill (63)

and naturally the forward (intrinsic) index spread as

$$S_f(t, T_e, T_m) = \frac{\sum_{j=1}^{J} w_j V_{prot,j}(t, T_e, T_m)}{\sum_{j=1}^{J} w_j A_j(t, T_e, T_m)}$$ \hfill (64)

which allows us to write the present value of the forward index as

$$V_f(t, T_e, T_m) = V_{prot,f}(t, T_e, T_m) - CA_f(t, T_e, T_m)$$

$$= (S_f(t, T_e, T_m) - C)A_f(t, T_e, T_m).$$ \hfill (65)
6.3 Default Adjusted Index Value

If the holder of a payer option on an index exercises that option at \( T_e \), they enter a payer CDS index at time \( T_e \), paying the coupon \( C \), until expiry at \( T_m \). In addition, any defaults between the start of the option and \( T_e \) are settled at \( t_{es} \). The clean value of the index plus the default settlement at \( t_{es} \) is

\[
V_{D,I}(T_e; T_m) = f(T_e) \left[ S_I(T_e, T_m) - C \right] \bar{A}_I(T_e, T_m | \mathcal{S}_I) + \sum_{j=1}^{J} \mathbb{1}_{t_j \leq T_e} w_j (1 - R_j) \tag{66}
\]

This is known as the default-adjusted forward portfolio swap price \([\text{Ped03}]\) or just the default-adjusted price. The full value of the index plus default settlement at \( T_e \) includes the accrued (on the remaining index) and the accrual-on-default (on the defaulted entries) so is given by

\[
V_{I,D}(T_e, T_m) = \Delta C \tag{67}
\]

6.3.1 Intrinsic Value

The default-adjusted price also has an intrinsic value given by

\[
V_{I,D}(T_e, T_m) = \sum_{j=1}^{J} w_j \left[ (1 - R_j) \mathbb{1}_{t_j \leq T_e} + \mathbb{1}_{t_j > T_e} (S_j(T_e, T_m) - C) A_j(T_e, T_m) \right] \tag{67}
\]

and its expectation under the \( \mathbb{T} \)-forward measure is

\[
F_{I,D}(t, T_e, T_m) = \mathbb{E}_{\mathbb{T}} \left[ V_{I,D}(T_e, T_m) \right] \equiv \mathbb{E}_{\mathbb{T}} \left[ V_{I,D}(T_e, T_m) \right] \tag{68}
\]

\[
= \sum_{j=1}^{J} w_j \mathbb{1}_{t_j \leq t} (1 - R_j) + \sum_{j=1}^{J} w_j \mathbb{1}_{t_j > t} (1 - R_j) (1 - Q_j(t, T_e)) + \frac{V_I(t, T_e, T_m)}{P(t, t_{es})}.
\]

We call this the \textit{ATM forward price}. The first term in the expression is just value from the names in index that have already defaulted by \( t \); the second term is the expected value from defaults between \( t \) and \( T_e \); and the final term is expected value of the index (equation 62). All these terms can be calculated from the (adjusted) credit curves of the constituent single-names.

At expiry, the ATM forward price is equal to the default-adjusted (index) price, i.e.

\[
F_{I,D}(T_e, T_e, T_m) = V_{I,D}(T_e, T_m).
\]

6.3.2 Homogeneous Pool Approximation

If the pool is completely homogeneous, meaning the recovery rates and credit curves of the constituents are all identical, then the expected default-adjusted price reduces to

\[
F_{I,D}(t, T_e, T_m) = (1 - f(t))(1 - R) + f(t) \left[ (1 - R) (1 - Q(t, T_e)) \right] + \frac{f(t)}{P(t, t_{es})} \left( V_{prot}(t, T_e, T_m) - CA(t, T_e, T_m) \right) \tag{69}
\]
The protection leg and annuity are calculated for a full index from the single credit curve. The single credit curve, the index curve, can be built from index quotes for different terms, and hence the value of $F^D(t, T_e, T_m)$ estimated without knowing anything about the constituents names. Even if the pool is highly heterogeneous, approximating it this way and ignoring the underlying names, can still produce fairly accurate results. We show examples of this when we look at Greeks and for comparisons with Bloomberg.

6.4 Option Payoff

The exercise price of the option is given (contractually) by

$$G(K) = (K - C)\bar{A}(T_e, T_m|K))$$

where $\bar{A}(T_e, T_m|K)$ is the flat annuity for a spread level of $K$. This is quoted clean; the actual amount paid at the exercise settlement date, $t_{ex}$, to exercise the option is $G(K) - \Delta C$.

The option payoff at $T_e$ may be written as

$$\text{option payoff} = P(T_e, t_{ex})\left[\frac{\chi[V^D(T_e, T_m) - G(K)]}{\beta_T(T_e)} + \frac{\Delta C}{G(K)}\right].$$

We see that the accrual terms cancel, but this is purely because the accrual is included in the (full) exercise price.\footnote{For an index quoted on spread, these should first be converted to upfront amounts before bootstrapping the credit curve.}

If there have been no defaults by the exercise date, then the payoff is exactly zero if the index spread equals the strike. Figure 2 shows the payoff against index spread for a payer and receiver option with different numbers of defaults at expiry.

Formally, the option value is

$$V_{I,\text{option}}(t, T_e, T_m) = E^T_t\left[\frac{P(T_e, t_{ex})}{\beta_T(T_e)}\left(\chi[V^D(T_e, T_m) - G(K)]\right)^+\right].$$

The payoff depends both on the index spread and the number of defaults at expiry, so in principle we need to jointly model all the constituent names of the index. Following the normal change of measure rules, we also have

$$V_{I,\text{option}}(t, T_e) = P(t, t_{ex})E^T_t\left[\left(\chi[V^D(T_e, T_m) - G(K)]\right)^+\right]$$

where $T$ is the measure associated with the numeraire $P(\cdot, t_{ex})$. The excise price, $G(K)$ is strictly not known advance as it depends (weakly) on the yield curve seen at $T_e$. However it is usual to set $G(K)$ to its expected value at $t$ since the variance from this is negligible next to the variance of $V^D_T$.

What is required is a model for the distribution of $V^D_T(T_e, T_m)$ under the $T$ forward measure, which preserves its expectation given in equation 68. If the distribution of $V^D_T(T_e, T_m)$ is given by $\rho^T(F)$ such that

$$F^D_T(t, T_e, T_m) = \int_{-\infty}^{\infty} F\rho^T(F)dF$$

\footnote{This is something that is simply not mentioned in the literature we have reviewed.}
then the option price is
\[
V_{I,\text{option}}(t, T_e) = P(t, t_{es}) \int_{-\infty}^{\infty} (\chi [F - G(K)])^+ \rho^T(F) dF
\] (75)

Up until this point the analysis has been exact, and any distribution that satisfies equation 74 will produce arbitrage free prices.

### 6.5 Put-Call relation

The value of a receiver option is
\[
V_{I,\text{Receiver}}(t, T_e, T_m) = P(t, t_{es}) \mathbb{E}^T_t [(G(K) - V^D_{I}(T_e, T_m))^+]
\] (76)

So the put-call relation is
\[
V_{I,\text{Payer}} - V_{I,\text{Receiver}} = P(t, t_{es}) \mathbb{E}^T_t [V^D_{I}(T_e, T_m) - G(K)]
\]
\[
= P(t, t_{es}) (F^D_{I}(t, T_e, T_m) - G(K))
\] (77)

So once the payer (receiver) price is calculated, the receiver (payer) price is available (almost) for free.

The ATM strike is defined as the strike level that makes the payer and receiver options have the same value; in this case this is given by
\[
G(K_{ATM}) = F^D_{I}(t, T_e, T_m)
\] (78)

which justified us calling \( F^D_{I}(t, T_e, T_m) \) the ATM forward value. This also allows us to define an ATM forward spread as the value of \( K_{ATM} \) that solves the above equation.
6.6 Default Adjusted Forward Index Spread

This is often referred to as just the forward index spread, however it should not be confused with either the forward index spread, \(S_I(t, T_e, T_m)\) or the ATM forward spread \(K_{ATM}(t, T_e, T_m)\) we defined earlier. It is defined as the ratio of the protection (including default settlement) to the forward annuity, so

\[
S^D_I(t, T_e, T_m) = \frac{P(t, t_{es}) \left( \sum_{j=1}^{J} w_j (1 - R_j)(1 - Q_j(t, T_e)) \right) + \sum_{j=1}^{J} w_j V_{prot,j}(t, T_e, T_m)}{\sum_{j=1}^{J} w_j A_j(t, T_e, T_m)}
\]

(79)

This of course is undefined in the armageddon scenario that the entire pool defaults by \(t\).

The default-adjusted forward spread is a martingale in the measure with the intrinsic annuity, \(A_I(t, T_e, T_m)\), as the numeraire.

The present value of the default-adjusted index may be written as

\[
V^D_I(t, T_e, T_m) = E_t \left[ P(t, t_{es}) \frac{(S^D_I(T_e, T_m) - C)A_I(T_e, T_m)}{\beta(T_e)} \right] = A_I(t, T_e, T_m) E_t \left[ (S^D_I(T_e, T_m) - C) \right] = A_I(t, T_e, T_m)(S^D_I(t, T_e, T_m) - C)
\]

(80)

This is a useful result for the approximate option pricing formula we discuss in section 6.7.2.

6.6.1 Homogenous Pool

For a homogenous pool, the default-adjusted forward index spread may be written as

\[
S^D_I(t, T_e, T_m) = \frac{P(t, t_{es}) \left( \frac{1-f(t)}{f(t)} (1 - R_j)(1 - Q_j(t, T_e)) \right) + V_{prot,I}(t, T_e, T_m)}{A_I(t, T_e, T_m)}
\]

(81)

6.7 Option Pricing Models

6.7.1 The Pedersen Model

The following strategy is due to Pedersen [Ped03] - it is very similar to what was done in section 4.2.2. Let the terminal distribution of \(V^P_I(T_e, T_m)\) under \(\mathbb{T}\) be a function of a pseudo index spread, \(\tilde{X}\) (call default-adjusted forward portfolio spread in [Ped03]), which absorbs the payments from the defaulted names.

\[
V^P_I(T_e, T_m) = (\tilde{X} - C)\tilde{A}(T_e, T_m|\tilde{X})
\]

\[
\tilde{X} = \tilde{X}_0 \exp \left( -\frac{\sigma^2}{2} (T_e - t) + \sigma Z \sqrt{T_e - t} \right)
\]

(82)

where \(Z\) is a standard Gaussian random variable, the annuity is calculated on the full index (i.e. no defaults, index factor, \(f = 1\)), and the value of \(\tilde{X}_0\) must be chosen so that equation 74 is satisfied.\(^{39}\)

\(^{39}\)A process for the default-adjusted spread, \(X_t\), must by definition jump at defaults, so a log-normal distribution for \(X_T\) may be lacking enough kurtosis.
We have found that as little as seven is sufficient.

An upper limit must be chosen; the option price (for a spread based on $Z$) is not well represented by a polynomial (it is discontinuous in its first derivative). However, here we cannot use Gauss-Hermite quadrature since the function multiplying the Gaussian density is not too onerous. Of course, $X_0$ depends on the volatility, $\sigma$, so must be recomputed for each value it takes.

The option value is now given by

$$V_{\text{option}}(t, T_e, T_m) = P(t, t_{es}) \int_{-\infty}^{\infty} \left( \chi \left[ V^D_I(T_e, T_m | \bar{X}_0, Z) - G(K) \right] \right)^+ \phi(Z) dZ$$

where $Z^*$ solves $V^D_I(T_e, T_m | \bar{X}_0, Z) = G(K)$. This integral must also be computed numerically, however here we cannot use Gauss-Hermite quadrature since the function multiplying the Gaussian density is not well represented by a polynomial (it is discontinuous in its first derivative). It is better to use the second form (after solving for $Z^*$) and compute it using Simpson’s rule (or some higher order scheme) [PTVF07].

As an example of this in practice, we show in figure 3 the option price (for a spread based strike) where between 0 and 3 names in the index have already defaulted. The option has one month to expiry, and the index (which is a randomly generated example) initially had 100 (equally weighted) names.

### 6.7.2 A Modified Black Formula

We may write the option price in terms of the default-adjusted forward index spread:

$$V_{\text{option}}(t, T_e) = \mathbb{E}_t \left[ \frac{P(T_e, t_{es})}{\beta_t(T_e)} \left[ \chi(\sqrt{[S^D_I(T_e, T_m) - C] A_I(T_e, T_m) - G(K)})^+ \right] \right]$$

$$= \frac{A_I(t, T_e, T_m) \mathbb{E}_t^A \left[ \chi([S^D_I(T_e, T_m) - \bar{K}]^+ \right]}{\mathbb{E}_t \left[ \frac{G(K)}{A_I(T_e, T_m)} \right]}$$

where $\bar{K} = C + \frac{G(K)}{A_I(T_e, T_m)}$

This expression is exact, but since the modified strike, $\bar{K}$ depends on the intrinsic index annuity at $T_e$, we cannot use this directly. However, if we make an approximation and set the modified
Figure 3: Payer option price when a number of names have defaulted from the index. The option is one month from expiry and the index initially had 100 names.

strike to be
\[ \tilde{K} = C + \frac{G(K)P(t, t_{cs})}{A_I(t, T_e, T_m)} \]  
which can be calculated at time \( t \), then since \( S_I^D \) is a Martingale in the annuity measure, the option can be priced with the Black formula. That is
\[ V_{option}(t, T_e) \approx A_I(t, T_e, T_m)\text{Black}(S_I^D(t, T_e, T_m), \tilde{K}, T_e - t, \tilde{\sigma}, \chi). \]  

The put-call relationship is given by
\[ V_{J,\text{Payer}} - V_{J,\text{Receiver}} = A_I(t, T_e, T_m) \left( S_I^D(t, T_e, T_m) - \tilde{K} \right) = A_I(t, T_e, T_m) \left( S_I^D(t, T_e, T_m) - C \right) - P(t, t_{cs})G(K) = V_I^D(t, T_e, T_m) - P(t, t_{cs})G(K) = P(t, t_{cs}) \left( F_I^D(t, T_e, T_m) - G(K) \right) \]  
so this modified Black formula does preserve the put-call relationship, which is critical for any option pricing model.

We should not expect that for the same volatility level, \( \sigma \), we obtain the same option prices from the Pedersen model as this modified Black formula. There are two reasons for this; firstly and most obviously, the modified Black formula is an approximation since we ‘freeze’ the modified strike at today’s value to make the expectation tractable; secondly, and more subtly, the volatility in Pedersen’s model is the log-normal volatility of a flat pseudo spread, while the volatility here is the log-normal volatility of the defaulted adjusted forward spread - these are not the same
thing, so given them equal volatility will not correspond to equal option prices. The clearest way to demonstrate this is to produce option prices for a range of strikes with the Pedersen model using a flat volatility, then find the implied volatility for the modified Black model - that is, the volatility parameter that must be input into equation 87 to produce the same option price. Figure 4 shows a plot of this implied volatility when the volatility of the flat pseudo spread is 40%. There are two options: one with one week and the other with three months to expiry. The two resultant volatility 'smiles' are extremely close.

![Modified Black Implied Volatility](image)

**Figure 4:** The implied volatility using the modified Black formula when option prices have been computed using the Pedersen model. The volatility used to generate the prices is 40%.

In this example the index coupon is 100bps, and when the strike is at this level, the exercise price, $G(K)$, is zero and the modified strike, $\tilde{K}$, is just $C$, so the modified Black formula is exact. We find the implied volatility for a strike of 100bps to be close to 40% (40.19% and 40.24% respectively), so the residual difference is because the two models use different spreads.

### 6.8 The Armageddon Scenario

The armageddon scenario is the theoretically possible, but unlikely, event that every name in an index defaults before the expiry of the option. It is claimed that this is incorrectly ignored in index option pricing. If the index annuity is used as a numeraire (as in the modified Black formula), then this can become zero while the option payoff is positive, which makes the option price theoretically unsound. However, while the default of every name means that forward spread given above is undefined, this has no effect on the pricing of index options we have presented using Pedersen’s model.
The value of $V^D_I(T_e, T_m)$ in the event of a total default of the index by $T_e$ is given by

$$V^D_I(T_e, T_m) = \sum_{j=1}^{J} w_j (1 - R_j)$$

which comes purely from the default settlement. The option payoff is

$$\text{armageddon option payoff} = P(T_e, t_{es}) \left( \chi \left[ \sum_{j=1}^{J} w_j (1 - R_j) - G(K) \right] \right)^+ \quad (89)$$

which is also well defined since the calculation of the exercise price, $G(K)$, does not depend on the number of names remaining in the index (even if this is zero). We price the option by taking the expectation under the $T$ forward measure which does not disappear in armageddon scenario. In fact the armageddon scenario simply represents the largest value $V^D_I(T_e, T_m)$ can take. Under Pedersen’s model, we have

$$\lim_{X \to \infty} (X - C) A(T_e, T_m | X) = 1 - R_I$$

so provided that $R_I \approx \sum_{j=1}^{J} w_j R_j$, the model included the armageddon scenarios in its states of the world.\(^{42}\)

Some attempts have been made to estimate the probability of armageddon from tranche prices \([BM09]\), but these do not exist for many indices, and where they do the results are somewhat suspect.

### 6.8.1 Extreme Strike Value

The limit of the exercise price as $K \to \infty$ is $1 - R_I$. It is often assumed that the payer option price will go to zero in this limit. However the realised value of $V^D_I(T_e, T_m)$ can exceed this limit (even without full default of the index) if there are enough defaults with realised recovery rates lower than expected (so that the default settlement is higher than expected), so the limit of the payer option price as $K \to \infty$ should be a (very) small positive number. Pedersen’s model does not capture this effect. In the modified Black case, the default adjusted forward spread, $S^D$, can go arbitrarily high, so the option will have some (very low) value even for $K \to \infty$.

### 7 Numerical Implementation

Key to a fast implementation of Pedersen’s model is the calculation of the flat annuity function $A(t, T | S)$. The market standard way to do this is the ISDA upfront model, in which an intermediary constant hazard rate is solved for (with a root finding algorithm), then this is used to compute the annuity. An often used approximation is to assume that the coupons on the premium are paid continually. For constant hazard rate pricing this gives the annuity as

$$A(t, T | \lambda) = \frac{\eta}{P(t, t_c)} \int_{t}^{T} P(t, u) e^{-\lambda u} dt \quad (90)$$

where $\eta$ is the ratio of the year fraction of the interval measured using the accrual day count convention, to the interval measured using the curve day count convention.\(^{43}\) With the annuity

\(^{42}\)We are not suggesting that the model gets the (risk neutral) probability of armageddon correct, just that it is an available state.

\(^{43}\)In most cases it will be simply 365/360 $\approx$ 1.0139.
calculated this way, the spread is given by

\[ S = \frac{(1 - R)\lambda}{\eta}. \]  

(91)

This gives us the hazard rate for a given spread level without root finding. The \( \eta \) term is normally missing in the credit triangle, however it greatly improves the estimate of \( \lambda \) at no real additional computational cost. The approximated hazard rate can then be used in equation 90 to give the annuity. Since the ISDA model uses log-linear interpolation for the discount factor \[ \text{[Whi13]}, \] the integral can be split up at the knots of the yield curve and computed as the sum of simple expressions; we call this approximation 1. We may also use the approximated hazard rate with a full calculation of the annuity (i.e. coupon paid every three months with accrual on default); we call this approximation 2.

Figure 5 shows the annuity for a 10Y CDS computed exactly and using the two credit triangle approximations. Approximation 2 is better for low spreads, while approximation 1 becomes better for high spreads. We prefer approximation 2, although the choice is not critical,\(^{44}\) as the effect of the calculation of the annuity on the final option price is small since much of the variation is absorbed in the calibration of \( X_0 \).

![Annuity Vs. Spread](attachment:Image.png)

Figure 5: The annuity for 10Y CDS computed exactly and using the two credit triangle approximations. The three curves cannot be distinguished visually, and the relative errors are shown on the right hand scale.

Figure 6 shows out-the-money (OTM) option prices against strike for a one month option. The index has 125 names (none defaulted) and the individual recovery rates and credit curves have been randomly generated. Shown is the option price using both the ISDA model and the credit triangle approximation to compute the annuity. The error is lowest for ATM options at

\(^{44}\)We could switch between the two approximations, but since this would mean a jump in annuity at the boundary we would need to split the integral in two to maintain high accuracy.
only 0.02%.

We find the time to price a single option (once the ATM forward price can been computed) to be 5.8ms using the full annuity calculation (ISDA upfront) and 1.4ms using the approximation.\footnote{These numbers apply to specific hardware so only the relative times are meaningful.}

Given that the error is so small, and time to price a single option is about four times faster, there is no compelling reason not to use the approximation. However, as the exercise prices are actual cash payments, these should be calculated exactly - this has no material effect on the time to price an option as it is a one-off calculation.

It is worth pointing out that the timing difference is as small as it is, because we have used our own, cleanroom, implementation of the ISDA model, which is extremely fast. This and all the code used to produce the results in this paper is available in the OpenGamma open source release.

\section{Greeks}

In this section we discuss the sensitivity of the price of an option (to enter a CDS index), to both market observables (state variables) and model parameters - these are collectively known as the Greeks.

We showed in \cite{Whi13} how to compute analytic Greeks for CDS prices, and this can be extended to compute analytic Greeks for options on CDS. However, here we restrict ourselves to the calculation of the Greeks by a \textit{bump and reprice method} - while this can be seen as just a forward finite difference approximation to the (mathematical) derivative $\frac{\partial V}{\partial X}$, where $V$ is the
option price and $X$ is the state variable or model parameter, it is these finite difference Greeks rather than the exact mathematical derivatives (analytic Greeks) that are usually given.

When we discuss individual Greeks below, we will present values for options to enter the CDS.NA.HY.21 5Y index. The trade date is 13-Feb-2014, the expiry is 19-Mar-2014 and the maturity (of the index) is 20-Dec-2018. The market date is given in table 10. We also use par spread data (from Markit) on the individual names in the index to build their credit curve. The examples we give use the Pedersen model to price the index options with inputs calculated using the ISDA model, however the technique is quite general and can be applied to any pricing model.

8.1 General Bump and Reprice

The Greek corresponding to the variable $X^{46}$ is usually calculated as

$$ G_X = \frac{V(X + \Delta X) - V(X)}{\Delta X} $$

where $\Delta X$ is the bump size - for some Greeks the difference is not divided by $\Delta X$. As we already mentioned, this is a poor estimate for the derivative $\frac{dV}{dX}$, but for historical reasons it is this number and not the derivative that is given.\(^{47}\)

Often we have computed the option price $V$ through some full pricing method, but wish, for speed, to calculate the bumped value $V(X + \Delta X)$ through some quicker (but less accurate) method.\(^{48}\) In this case it is important to also calculate the base price, $V(X)$, with this quick method rather than using the full computed value - this will lead to cancellation of errors that will give a value of the Greek closer to the value compute using the full model for the bumped price.

8.2 CS01 or Spread DV01

This is normally defined as the change in value of option due to a 1bps increase in spread. There are several ways this could be computed.

The simplest (and quickest to compute) method is to assume a homogenous pool with a single constant hazard rate calibrated to the price or spread of the underlying index. The index spread is then bumped up by 1bps and a new (bumped) hazard rate computed. The difference in the option price with these two hazard rates is the CS01 - we call this the flat spread method. It is very similar to what is done to calculate the CS01 for an index or single-name CDS (priced with a constant hazard rate) \(^{[Whi13]}\).

Another method which also assumes an homogenous pool, involves building an index credit curve from the market quotes of the index at several terms. One then simply bumps the spread quotes\(^{49}\) up by 1bps, rebuilds a bumped index credit curve and uses this to compute a bumped option price. The difference from the base price (i.e. the price calculated with the original index curve) is the CS01. We call this the index curve method.

If we are pricing considering the individual credit curves (adjusted to match index prices as in section 5.5), then one approach is to bump the (implied) spreads of all the curves by 1 bps, then

\[^{46}\]This can be a scalar or vector.

\[^{47}\]If one were estimating the derivative, one would use the central difference $\frac{dV}{dX} \approx \frac{V(X+\Delta X) - V(X-\Delta X)}{2\Delta X}$ which has an accuracy of $O(\Delta X^2)$ rather than the one sided difference which only has an accuracy of $O(\Delta X)$ - see any text on numerical analysis for details.

\[^{48}\]In practice there may be hundreds of ‘Greeks’ to calculate.

\[^{49}\]If the quotes are given as PUF, they must first be converted to quoted spreads.
reprice the option and take the difference. The trouble with this approach is that a 1bps shift is applied to all curves regardless of their current level, so high credit quality entries will have a relatively large shift compared to the lower credit quality entries in the index. An alternative is to first bump the quoted spread of the underlying index by 1bps, then readjust the individual curves to match this price. The option price with these readjusted curves is the bumped value. We call this the intrinsic method.

There are of course other ways to bump spreads that will give some measure of CS01. However, we restrict ourselves to the three methods presented above. The results are shown in table 3. The greatest variation between the methods is for ATM options, but even there the largest difference is less than $10 (on a notional of $100MM). In section 9 we compare the CS01 numbers from Bloomberg for the same example.

<table>
<thead>
<tr>
<th>Strike Price</th>
<th>Flat Index Spread</th>
<th>Index Curve</th>
<th>Intrinsic</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Payer</td>
<td>Receiver</td>
<td>Payer</td>
</tr>
<tr>
<td>103</td>
<td>138.95</td>
<td>-45379.86</td>
<td>139.05</td>
</tr>
<tr>
<td>104</td>
<td>781.47</td>
<td>-44737.34</td>
<td>781.92</td>
</tr>
<tr>
<td>105</td>
<td>3269.04</td>
<td>-42249.77</td>
<td>3270.49</td>
</tr>
<tr>
<td>106</td>
<td>9939.58</td>
<td>-35579.22</td>
<td>9942.71</td>
</tr>
<tr>
<td>107.144443</td>
<td>23642.68</td>
<td>-21876.12</td>
<td>23646.91</td>
</tr>
<tr>
<td>108</td>
<td>34507.50</td>
<td>-11011.31</td>
<td>34510.81</td>
</tr>
<tr>
<td>109</td>
<td>42497.16</td>
<td>-3021.64</td>
<td>42498.52</td>
</tr>
<tr>
<td>110</td>
<td>45100.02</td>
<td>-418.79</td>
<td>45100.28</td>
</tr>
</tbody>
</table>

Table 3: The CS01 for options to enter the CDX.NA.HY.21 5Y index calculated using the three methods discussed in the main text.

8.2.1 Bucketed CS01

A full bucketed CS01 would involve bumping each par spread50 of each individual underlying credit curve in turn and computing an option price. This will produce several hundred numbers (a CS01 vector). One use for this very granular approach is a delta based VaR calculation, where scenarios of spread changes are known (e.g. from historical data). These delta spreads can be multiplied by the CS01 vector to produce a change in option price for that scenario. Further discussion is outside the scope of this paper.

Another form of bucketed CS01 is sensitivity to points along the index credit curve. One way to proceed is to build an index curve from market quotes of the index at different terms. This curve can then be used to imply index par spreads for different times (the buckets), not necessarily corresponding to the index terms. By bumping these (implied) spreads in turn we can build new index curves, then use these directly to price the option (homogenous pool approximation) or first adjust the underlying credit curves to the index prices implied by the new (bumped) index curve. Either way we will obtain a set of several CS01 corresponding to the buckets we specified.

50Since we usually first adjust the curve to match index prices, there are the adjusted spreads implied from the adjusted credit curves.
8.3 Delta
An option Delta is usually defined as the sensitivity of the option price to the price of the underlying - i.e. the first derivative of the option price with respect to the price of the underlying. If the underlying is taken as the spot index price, this gives a spot delta, while if is it taken as the default adjusted forward index, this gives a forward delta.

8.3.1 Forward Delta
We define this as the sensitivity to the ATM forward price

\[ \Delta_F(t, T_e, T_m) = \frac{\partial V_{option}(t, T_e, T_m)}{\partial F_D(t, T_e, T_m)} \]  

As usual, this can approximated by finite difference.

8.3.2 Spot Delta
Delta in the context of index options should be the sensitivity of the option price to the price of the underlying index. If the price of the index is bumped up by a small amount, then we can readjust the individual credit curves to match this new index price, then price the option from these curves. This is remarkably similar to the calculation of CS01 where we bumped the index spread rather than the index price. This leads to a natural definition of the delta as

\[ \Delta(t, T_e, T_m) = \frac{CS01_{option}(t, T_e, T_m)}{CS01_I(t, T_m)} \]  

This is a hedge ratio, since if I have a portfolio of one option and \(-\Delta\) of the index, the portfolio’s CS01 will be zero (it is important that both CS01s are measuring the same spread sensitivity). It is this CS01 ratio that Bloomberg use to compute the option delta. The measure is consistent with the analytic delta since,

\[ \Delta(t, T_e, T_m) = \frac{\partial V_{option}(t, T_e, T_m)}{\partial V_I(t, T_m)} = \frac{\partial^2 V_{option}(t, T_e, T_m)}{\partial S^2} \frac{\partial S}{\partial V_I(t, T_m)} \]  

To show this in practice we compute the delta for our example options. The two methods are central finite difference (where the underlying credit curves are readjusted to the bumped index prices) and the CS01 ratio. In the second case the index CS01 is computed by bumping the index quoted spread by one basis point (this is the standard way to compute an index CS01), while the option CS01 is computed by our flat index spread method. Table 4 shows the computed delta for payer and receiver options with these two methods. Again the numbers are close, but they are not expected to be exact as they are different things: one is an approximation to the mathematical delta, while the other is the hedge ratio to make a portfolio insensitive to a one basis point rise in the quoted spread of the index.

Finally in figure 7 we show a plot of the delta (computed with central finite difference), which shows a pretty typical shape.

8.4 Gamma
Gamma is the only second order Greek we will consider. Formally (spot) Gamma is defined as

\[ \Gamma(t, T_e, T_m) = \frac{\partial^2 V_{option}(t, T_e, T_m)}{\partial V^2_F(t, T_m)} = \frac{\partial \Delta(t, T_e, T_m)}{\partial V_I(t, T_m)} \]  

32
<table>
<thead>
<tr>
<th>Strike Price</th>
<th>Payer</th>
<th>Receiver</th>
<th>Payer</th>
<th>Receiver</th>
</tr>
</thead>
<tbody>
<tr>
<td>103</td>
<td>0.29%</td>
<td>-99.70%</td>
<td>0.31%</td>
<td>-99.69%</td>
</tr>
<tr>
<td>104</td>
<td>1.64%</td>
<td>-98.34%</td>
<td>1.72%</td>
<td>-98.27%</td>
</tr>
<tr>
<td>105</td>
<td>6.95%</td>
<td>-93.04%</td>
<td>7.18%</td>
<td>-92.81%</td>
</tr>
<tr>
<td>106</td>
<td>21.33%</td>
<td>-78.66%</td>
<td>21.83%</td>
<td>-78.16%</td>
</tr>
<tr>
<td>107.14444</td>
<td>51.25%</td>
<td>-48.74%</td>
<td>51.94%</td>
<td>-48.06%</td>
</tr>
<tr>
<td>108</td>
<td>75.26%</td>
<td>-24.72%</td>
<td>75.80%</td>
<td>-24.19%</td>
</tr>
<tr>
<td>109</td>
<td>93.13%</td>
<td>-6.86%</td>
<td>93.35%</td>
<td>-6.64%</td>
</tr>
<tr>
<td>110</td>
<td>99.03%</td>
<td>-0.96%</td>
<td>99.07%</td>
<td>-0.92%</td>
</tr>
<tr>
<td>111</td>
<td>99.93%</td>
<td>-0.06%</td>
<td>99.94%</td>
<td>-0.05%</td>
</tr>
</tbody>
</table>

Table 4: The spot delta for options on CDX.NA.HY using the two methods discussed in the main text.

Figure 7: The delta for a payer option on the CDX.NA.HY.21 5Y index. The index is priced based, so the strike is given as an index price.
The standard finite difference way to approximate this is

\[
\Gamma(t, T_e, T_m) \approx \frac{V_{\text{option}}(V_I + \Delta V_I) + V_{\text{option}}(V_I - \Delta V_I) - 2V_{\text{option}}(V_I)}{\Delta V_I} \tag{96}
\]

where \(V_{\text{option}}(X)\) means the option price computed with an index value of \(X\) using one of the methods we have already discussed. This can also be expressed (less accurately) as a forward finite difference using the Delta, i.e.

\[
\Gamma(t, T_e, T_m) \approx \frac{\Delta(V_I + \Delta V_I) - \Delta(V_I)}{\Delta V_I} \tag{97}
\]

### 8.4.1 Alternative Definition of Gamma

Bloomberg defines Gamma as the change in Delta for a 10bp rise in spread. Since Gamma may be computed as

\[
\Gamma(t, T_e, T_m) = \frac{\partial \Delta}{\partial S} \frac{\partial S}{\partial V_I} \approx (\Delta(S + 10bp) - \Delta(S)) \times 1000 \frac{\partial S}{\partial V_I} \tag{98}
\]

it can be viewed as a scaled version of Gamma. In terms of CS01 ratios this can be written as

\[
\Gamma_{\text{Bloomberg}} = \frac{\text{CS01}_{\text{option}}(S + 10bp)}{\text{CS01}_{\text{index}}(S + 10bp)} - \frac{\text{CS01}_{\text{option}}(S)}{\text{CS01}_{\text{index}}(S)}. \tag{99}
\]

In table 5 we show the Gamma calculated by finite difference and as the change in delta for a 10bp rise in the spread.

<table>
<thead>
<tr>
<th>Strike Price</th>
<th>103</th>
<th>104</th>
<th>105</th>
<th>106</th>
<th>107.14444</th>
<th>108</th>
<th>109</th>
<th>110</th>
<th>111</th>
</tr>
</thead>
<tbody>
<tr>
<td>Payer</td>
<td>0.6589</td>
<td>3.0471</td>
<td>9.9261</td>
<td>21.6243</td>
<td>29.6437</td>
<td>23.4869</td>
<td>9.8287</td>
<td>1.9173</td>
<td>0.1491</td>
</tr>
<tr>
<td>Receiver</td>
<td>0.6588</td>
<td>3.0471</td>
<td>9.9260</td>
<td>21.6242</td>
<td>29.6437</td>
<td>23.4868</td>
<td>9.8286</td>
<td>1.9173</td>
<td>0.1490</td>
</tr>
<tr>
<td>Payer</td>
<td>0.4941%</td>
<td>2.0107%</td>
<td>5.7525%</td>
<td>10.9853%</td>
<td>12.9199%</td>
<td>9.1124%</td>
<td>3.3217%</td>
<td>0.5631%</td>
<td>0.0379%</td>
</tr>
<tr>
<td>Receiver</td>
<td>0.4941%</td>
<td>2.0107%</td>
<td>5.7526%</td>
<td>10.9854%</td>
<td>12.9200%</td>
<td>9.1124%</td>
<td>3.3218%</td>
<td>0.5632%</td>
<td>0.0380%</td>
</tr>
</tbody>
</table>

Table 5: Two different definitions of Gamma: The mathematical definition (computed by finite difference) and the definition as the change in delta for a 10bps rise in the spread. The second definition is used by Bloomberg.

### 8.5 Vega

For Pedersen model, this has a straight forward interpretation. It is simply

\[
\nu(t, T_e, T_m) = \frac{\partial V_{\text{option}}(t, T_e, T_m)}{\partial \sigma} \tag{100}
\]
This is often calculated as the change in option price for a 1% point rise in spread volatility, i.e.
\[ \nu(t, T_e, T_m) \approx V_{option}(\sigma + 1\%) - V_{option}(\sigma). \]  
(101)

The Greeks we have calculated so far have involved bumping credit curves (either the intrinsic curves or an index curve), which affects the value of the ATM forward price, which in turn affects the option price. For Vega, the ATM forward price remains unchanged. This means that once an ATM forward price is established, the details of that calculation are irrelevant in the calculation of Vega.

In figure 8 we show a plot of Vega calculated using central finite difference (with a bump of 1bp which has been scaled to a 1% Vega). The value of Vega should be identical for payer and receiver options, and we find this is so (up to some numerical noise).

![Vega graph](image)

Figure 8: The Vega for an option on the CDX.NA.HY.21 5Y index. The index is priced based, so the strike is given as an index price.

### 8.6 Theta

Theta is the sensitivity of the option price to the time-to-expiry, so measures the time decay in the value of an option. It is often calculated as a one-day-theta - the difference in the option price one day ahead from the price now. Depending on what is held constant with respect to the time-to-expiry, different Theta can be calculated.

Since the ATM Forward value, \( F_D(t, T_e, T_m) \) is a Martingale (under the \( T \) forward measure), its expected value one day ahead should be just today’s value - so, just like for Vega, all we need to consider is the sensitivity of the option price to a change of time-to-expiry for a fixed ATM forward price. For a payer (receiver) option the potential for upside (downside) gains from taking large (small) values of the default-adjusted index value is reduced as the option duration

---

\(^{51}\)Of course this is just the forward finite difference approximation scaled down by 10,000.
reduces, so Theta is negative. This calculation implicitly assumes defaults can occur over the one day period (that is what makes $F^D_P(t, T_e, T_m)$ a Martingale).

An alternative (and unconventional) measure of Theta is the value conditional the credit curves remaining the same and no defaults occurring (over the one day period). Viewed this way, Theta consists of two main parts; the spread diffusion part and the default part. For payer options the potential for upside gains from taking large values of the index spread is reduced as the option duration reduces; and the potential for gains from defaults from the index (and their subsequent settlement) is reduced as the time over which defaults can occur is reduced - so Theta is always negative. For receiver options the potential for downside gains from taking small values of the index spread is reduced as the option duration reduces; but the potential for losses from defaults from the index is also reduced as the time over which defaults can occur is reduced - this can give a positive Theta for deeply in the money receiver options.

The ATM forward value is calculated one day ahead assuming the realised value of the curves (credit and yield) one day forward are just their expected value (and of course the index factor remains the same). Once the new ATM forward is calculated Theta is computed as above. These two definitions of Theta are shown for our example option in table 6.

For payer options the Theta is larger in magnitude (more negative) for the no defaults version of Theta - some of the option value comes from the default settlement, so not considering defaults over the one day period lowers the option value more; this is most striking for deeply in the money options. The opposite is true for receiver options, where deeply in the money options gain value (positive Theta).

<table>
<thead>
<tr>
<th>Strike Price</th>
<th>Payer</th>
<th>Receiver</th>
<th>Payer</th>
<th>Receiver</th>
</tr>
</thead>
<tbody>
<tr>
<td>103</td>
<td>-166.32</td>
<td>-148.17</td>
<td>-171.78</td>
<td>1998.55</td>
</tr>
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<td>104</td>
<td>-787.08</td>
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<td>-819.59</td>
<td>1346.36</td>
</tr>
<tr>
<td>105</td>
<td>-2610.19</td>
<td>-2600.78</td>
<td>-2752.95</td>
<td>591.37</td>
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<tr>
<td>106</td>
<td>-5754.79</td>
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</tr>
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<td>107.144434</td>
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<td>-7928.51</td>
<td>-9030.51</td>
<td>-6878.31</td>
</tr>
<tr>
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</tr>
<tr>
<td>109</td>
<td>-2580.19</td>
<td>-2588.28</td>
<td>-4590.44</td>
<td>-2446.36</td>
</tr>
<tr>
<td>110</td>
<td>-480.55</td>
<td>-493.02</td>
<td>-2613.78</td>
<td>-474.07</td>
</tr>
<tr>
<td>111</td>
<td>-20.09</td>
<td>-36.93</td>
<td>-2171.23</td>
<td>-35.89</td>
</tr>
</tbody>
</table>

Table 6: Theta calculated allowing and not allowing defaults over the one day period.

### 8.7 Interest Rate DV01

This is the sensitivity of the option price to the interest rate. It is normally calculated by calibrating a bumped yield curve with instruments whose rates have been bumped up by 1bp.\[52\] The IR DV01 of an index is defined as the price calculated with this bumped yield curve when the index spread is kept constant, minus the price with the original yield curve (and of course the same spread).

\[52\]In practice you will obtain a similar result by bumping the knots of the calibrated yield curve by 1bp.
We may calculate IR DV01 of an option by first adjusting the underlying credit curves to match the index spreads using the bumped yield curve. If we price the option with these curves, this gives us a yield curve bumped price. We will obtain very similar results using the homogenous pool approximation. Table 7 shows the IR DV01 for our example options. The two calculation methods show remarkable agreement (less than a $1 difference for ATM options on a notional of $100MM).

<table>
<thead>
<tr>
<th>Strike Price</th>
<th>Flat Index Spread</th>
<th>Intrinsic</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Payer</td>
<td>Receiver</td>
</tr>
<tr>
<td>103</td>
<td>1.53</td>
<td>-1800.47</td>
</tr>
<tr>
<td>104</td>
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<td>-1779.14</td>
</tr>
<tr>
<td>105</td>
<td>65.56</td>
<td>-1714.78</td>
</tr>
<tr>
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<td>247.55</td>
<td>-1521.96</td>
</tr>
<tr>
<td>107.1444</td>
<td>716.91</td>
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</tr>
<tr>
<td>111</td>
<td>1713.34</td>
<td>-2.03</td>
</tr>
</tbody>
</table>

Table 7: The IR DV01 for options on CDX.NA.HY calculated from a flat index curve and using the intrinsic credit curve.

It is well noted that the IR01 for CDS is one or two orders of magnitude smaller than the equivalent CS01 [O’k08, Whi13], and we see the same thing for these risk factors on options on CDS indices.

### 8.8 Default Sensitivity

This is the change in the option premium due to the default of a single-name from the underlying index. Unlike other Greeks, which are (mathematical) derivatives, this is a jump caused by a sudden unpredictable event. To calculate this value for a given name, we remove that name from the index and recalculate the ATM forward price assuming the remaining intrinsic credit curves are unaffected by the default.\(^{53}\) This ‘bumped’ ATM forward price is then used to recalculate the option premium and the difference is the default sensitivity (to that name).

We consider the same four names we looked at in section 5.5.2 (table 2), since they represent very different credit qualities in the index. Figure 9 shows the result of the calculation for a payer option. For deep in the money options (which are certain to be exercised), the default of AES Corp causes a jump in premium of 7.67 \(\times\) 10\(^{-3}\) - this is larger than the default settlement of 6 \(\times\) 10\(^{-3}\) (the recovery rate of AES is 40%). Since the credit quality of AES is higher than the index average, its removal raises the value of the index (to a payer of coupons), which accounts for the extra premium gain.\(^{54}\) At the opposite end, the default of TX Competitive Elec Hldgs causes a jump of only 2.72 \(\times\) 10\(^{-4}\), despite the fact that its default settlement is 9.4 \(\times\) 10\(^{-3}\) (it has a recovery rate of only 6%) - this is because the default is already figured in the index value (the spreads are massive), so the default settlement is balanced by a drop in the index value.

\(^{53}\)In reality, a default will lead to the widening of spreads of other names in the sector.

\(^{54}\)The index has a negative value, so removal of an ‘average’ name will increase its value (move it closer to zero), but removal of a better than average name will cause a larger change.
Figure 9: The default sensitivity for a option on the CDX.NA.HY.21 5Y index, for the default of specified names. The index is priced based, so the strike is given as an index price.

One could carry out this calculation for every name in the index. An alternative is to use the homogenous pool approximation and consider the loss of a single-name. After performing this calculation we obtain a curve that plateaus at $7.95 \times 10^{-3}$, ignoring discounting, this consists of $7.0 \times 10^{-3}$ from the default settlement (recall the index recovery rate is 30%) and the remaining from the change to the forward index value.\footnote{The PUF of the index will remain unchanged, so its value (per unit notional) will increase by $7.62 \times 10^{-4}$, i.e. -1\% of the PUF.}

**9 Comparison With Bloomberg**

In this section we give a few examples of single-name and index options taken from Bloomberg’s CDSO screen\footnote{This allows you to see calculated values of the option premium and Greeks for a user input volatility level. It does not give actual traded option prices.}, and show how their prices and Greeks compare with our own analytics.

**9.1 Single Name Default Swaptions**

Our first example is the option to enter the 5Y CDS referencing the republic of Italy. The trade date is 5-Feb-2014, the option expiry date is 20-Mar-2014 and the maturity of the underlying CDS is 20-Jun-2019. If the holder of a payer swaption exercises the option (on 20-Mar-2014) they enter a (bespoke) CDS with a running coupon equal to the strike and accrual starting from 21-Mar-2014 (i.e. T+1). This type of option was discussed in section 4.2.1 and can be priced using the Black formula.
The yield curve is the ISDA Standard USD snapped at 1600 (New York time) on the 4-Feb-2014 using instruments (deposits and swaps) with spot date of 7-Feb-2014. The credit curve is built using spreads\(^{57}\) and standard CDS. This market data is given in table 8.

<table>
<thead>
<tr>
<th>Period</th>
<th>Type</th>
<th>Rate</th>
</tr>
</thead>
<tbody>
<tr>
<td>1M</td>
<td>Deposit</td>
<td>0.16%</td>
</tr>
<tr>
<td>2M</td>
<td>Deposit</td>
<td>0.20%</td>
</tr>
<tr>
<td>3M</td>
<td>Deposit</td>
<td>0.24%</td>
</tr>
<tr>
<td>6M</td>
<td>Deposit</td>
<td>0.33%</td>
</tr>
<tr>
<td>1Y</td>
<td>Deposit</td>
<td>0.56%</td>
</tr>
<tr>
<td>2Y</td>
<td>Swap</td>
<td>0.44%</td>
</tr>
<tr>
<td>3Y</td>
<td>Swap</td>
<td>0.78%</td>
</tr>
<tr>
<td>4Y</td>
<td>Swap</td>
<td>1.19%</td>
</tr>
<tr>
<td>5Y</td>
<td>Swap</td>
<td>1.58%</td>
</tr>
<tr>
<td>6Y</td>
<td>Swap</td>
<td>1.92%</td>
</tr>
<tr>
<td>7Y</td>
<td>Swap</td>
<td>2.19%</td>
</tr>
<tr>
<td>8Y</td>
<td>Swap</td>
<td>2.42%</td>
</tr>
<tr>
<td>9Y</td>
<td>Swap</td>
<td>2.61%</td>
</tr>
<tr>
<td>10Y</td>
<td>Swap</td>
<td>2.76%</td>
</tr>
<tr>
<td>12Y</td>
<td>Swap</td>
<td>3.01%</td>
</tr>
<tr>
<td>15Y</td>
<td>Swap</td>
<td>3.25%</td>
</tr>
<tr>
<td>20Y</td>
<td>Swap</td>
<td>3.46%</td>
</tr>
<tr>
<td>25Y</td>
<td>Swap</td>
<td>3.55%</td>
</tr>
<tr>
<td>30Y</td>
<td>Swap</td>
<td>3.59%</td>
</tr>
</tbody>
</table>

<table>
<thead>
<tr>
<th>Tenor</th>
<th>Spread (bps)</th>
</tr>
</thead>
<tbody>
<tr>
<td>6M</td>
<td>57.4300</td>
</tr>
<tr>
<td>1Y</td>
<td>74.9700</td>
</tr>
<tr>
<td>2Y</td>
<td>111.3200</td>
</tr>
<tr>
<td>3Y</td>
<td>139.3200</td>
</tr>
<tr>
<td>4Y</td>
<td>157.6400</td>
</tr>
<tr>
<td>5Y</td>
<td>173.6600</td>
</tr>
<tr>
<td>7Y</td>
<td>209.2800</td>
</tr>
<tr>
<td>10Y</td>
<td>228.3500</td>
</tr>
</tbody>
</table>

<table>
<thead>
<tr>
<th>Interest Rate Conventions</th>
</tr>
</thead>
<tbody>
<tr>
<td>Spot Date</td>
</tr>
<tr>
<td>Swap DCC</td>
</tr>
<tr>
<td>MM DCC</td>
</tr>
<tr>
<td>Swap Interval</td>
</tr>
<tr>
<td>Floating DCC</td>
</tr>
<tr>
<td>Holidays</td>
</tr>
<tr>
<td>Floating Interval</td>
</tr>
<tr>
<td>Bad Day Conv</td>
</tr>
</tbody>
</table>

Table 8: Market data used to build the Standard ISDA USD yield curve (left), and Spreads on Republic of Italy CDSs for 5-Feb-2014 (right).

Having calibrated the yield and credit curve we compute the forward spread using equation 23, where the (forward) protection leg and annuity are computed using the ISDA model. We obtain a value of 182.764 bps compared to Bloomberg’s value (ATM Fwd) of 182.767 bps. The annuity we calculate as 4.8354\(^{58}\), which compares to the Bloomberg implied value\(^{59}\) of 4.8364.

\(^{57}\)These are treated as par spreads, even though the numbers given are actually the quoted spreads for those maturities.

\(^{58}\)For unit notional and per unit of spread.

\(^{59}\)This value is not given on the CDSO screen, we have implied it by looking at call-put values across a range of strikes.
We consider a range of strikes from 100 to 300 bps with a spread volatility of 40%. The time to expiry is taken as 0.11781 years. In table 9 we show the calculated option premium and the implied volatility from the premiums given on the CDSO screen. Black1 uses our calculated values for the forward spread and annuity, while Black2 uses the Bloomberg values for these qualities. The implied volatility is the number that must be put into the Black formula to recover the premium given by CSDO (again Imp Vol1 and Imp Vol2 use our and Bloomberg’s forward spread and annuity). In both cases the implied volatility is around 39% rather than the 40% we would expect for options priced with the Black formula. The discrepancy is much large than any error from calculating the forward spread and annuity, or from the calculation of the time-to-expiry. The only explanation we can offer is than Bloomberg is using some other model to price these type of options.

<table>
<thead>
<tr>
<th>Strike</th>
<th>Option Premium</th>
<th>Implied Volatility</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Bloomberg</td>
<td>Black1</td>
</tr>
<tr>
<td>100.000</td>
<td>4.0029595%</td>
<td>4.0020013%</td>
</tr>
<tr>
<td>140.000</td>
<td>2.0777336%</td>
<td>2.0783381%</td>
</tr>
<tr>
<td>150.000</td>
<td>1.6186780%</td>
<td>1.6211885%</td>
</tr>
<tr>
<td>160.000</td>
<td>1.1956130%</td>
<td>1.2009104%</td>
</tr>
<tr>
<td>170.000</td>
<td>0.8304912%</td>
<td>0.8386714%</td>
</tr>
<tr>
<td>180.000</td>
<td>0.5396918%</td>
<td>0.5497805%</td>
</tr>
<tr>
<td>182.767</td>
<td>0.4732540%</td>
<td>0.4835945%</td>
</tr>
<tr>
<td>190.000</td>
<td>0.3274686%</td>
<td>0.3378529%</td>
</tr>
<tr>
<td>200.000</td>
<td>0.1856680%</td>
<td>0.1948579%</td>
</tr>
<tr>
<td>210.000</td>
<td>0.0986346%</td>
<td>0.1057882%</td>
</tr>
<tr>
<td>220.000</td>
<td>0.0492900%</td>
<td>0.0542785%</td>
</tr>
<tr>
<td>230.000</td>
<td>0.0232763%</td>
<td>0.0264402%</td>
</tr>
<tr>
<td>250.000</td>
<td>0.0044651%</td>
<td>0.0054703%</td>
</tr>
<tr>
<td>300.000</td>
<td>0.0000367%</td>
<td>0.0000583%</td>
</tr>
</tbody>
</table>

Table 9: Comparison of option premiums given on CDSO to that calculated from the Black formula. The spread volatility is 40%, and implied volatility is for the premium quoted on CDSO. The sets of data 1 & 2 refer to what value is used for the forward and numeraire in the Black formula. See main text for details.

9.2 Index Options - CDX High Yield

Our first index option example is an option to enter the 5Y CDX.NA.HY.21-V1 - this is the same as was used in the Greeks section. The reason we have chosen to start with this index is because it trades on price rather than spread, so there are no issues with converting from a strike given as a spread to an exercise price. The notional amount is $100MM, the trade date is 13-Feb-2014, the option expiry is 19-Feb-2014, and the maturity of the underlying index is 20-Dec-2018. The index has a coupon of 500bps and a recovery rate of 30%.

The current price is 107.62 (so PUF of -7.62) which corresponds to a quoted spread of 322.5621

---

60 Using ACT/ACT ISDA day count between 5-Feb-2104 and 20-Mar-2014.
61 This is chosen so there is no accrued premium.
bps.\textsuperscript{62} The market data used to build the yield curve (ISDA Standard USD) and the prices for other terms of CDS.NA.HY.21 are given in table 10. It is worth noting that the CDSO screen gives prices at the standard tenors of 6M, 1Y, 2Y, 3Y, 4Y, 5Y, 7Y and 10Y even though only 3Y, 5Y, 7Y and 10Y trade on-the-run.\textsuperscript{63} The prices for tenors less than 3Y are set to the 3Y value, but only the 3Y, 5Y, 7Y and 10Y appear to be used by Bloomberg for curve construction. We also only use these points, and that is what we show in the table. Of course only the 3Y and 5Y points can affect the value of options on the 5Y index.

<table>
<thead>
<tr>
<th>Period</th>
<th>Type</th>
<th>Rate</th>
</tr>
</thead>
<tbody>
<tr>
<td>1M</td>
<td>Deposit</td>
<td>0.15%</td>
</tr>
<tr>
<td>2M</td>
<td>Deposit</td>
<td>0.20%</td>
</tr>
<tr>
<td>3M</td>
<td>Deposit</td>
<td>0.24%</td>
</tr>
<tr>
<td>6M</td>
<td>Deposit</td>
<td>0.33%</td>
</tr>
<tr>
<td>1Y</td>
<td>Deposit</td>
<td>0.55%</td>
</tr>
<tr>
<td>2Y</td>
<td>Swap</td>
<td>0.47%</td>
</tr>
<tr>
<td>3Y</td>
<td>Swap</td>
<td>0.83%</td>
</tr>
<tr>
<td>4Y</td>
<td>Swap</td>
<td>1.27%</td>
</tr>
<tr>
<td>5Y</td>
<td>Swap</td>
<td>1.67%</td>
</tr>
<tr>
<td>6Y</td>
<td>Swap</td>
<td>2.02%</td>
</tr>
<tr>
<td>7Y</td>
<td>Swap</td>
<td>2.30%</td>
</tr>
<tr>
<td>8Y</td>
<td>Swap</td>
<td>2.53%</td>
</tr>
<tr>
<td>9Y</td>
<td>Swap</td>
<td>2.72%</td>
</tr>
<tr>
<td>10Y</td>
<td>Swap</td>
<td>2.89%</td>
</tr>
<tr>
<td>12Y</td>
<td>Swap</td>
<td>3.14%</td>
</tr>
<tr>
<td>15Y</td>
<td>Swap</td>
<td>3.38%</td>
</tr>
<tr>
<td>20Y</td>
<td>Swap</td>
<td>3.58%</td>
</tr>
<tr>
<td>25Y</td>
<td>Swap</td>
<td>3.67%</td>
</tr>
<tr>
<td>30Y</td>
<td>Swap</td>
<td>3.71%</td>
</tr>
</tbody>
</table>

<table>
<thead>
<tr>
<th>Tenor</th>
<th>Price</th>
</tr>
</thead>
<tbody>
<tr>
<td>3Y</td>
<td>107.5600</td>
</tr>
<tr>
<td>5Y</td>
<td>107.6200</td>
</tr>
<tr>
<td>7Y</td>
<td>105.7100</td>
</tr>
<tr>
<td>10Y</td>
<td>106.5200</td>
</tr>
</tbody>
</table>

CDS Conventions

| Frequency | Quarterly IMM |
| Day Count | ACT/360 |
| Recovery Rate | 30% |

Interest Rate Conventions

| Spot Date | 17-Feb-2014 |
| Swap DCC | 30/360 |
| MM DCC | A/360 |
| Swap Interval | 6M |
| Floating DCC | A/360 |
| Holidays | none |
| Floating Interval | 3M |
| Bad Day Conv | MF |

Table 10: Market data used to build the Standard ISDA USD yield curve (left), and prices of for CDX.NA.HY.21-V1 on 13-Feb-2014 (right).

\textsuperscript{62}Recall the conversion is via the ISDA model.

\textsuperscript{63}Seasoned indices (off-the-run) exist that may have, say, 6 months to maturity, but these are highly illiquid and generally will have different constituents to the on-the-run indices.
9.2.1 ATM Forward

What Bloomberg calls the ATM Forward for price based options, when expressed as an exercise price, gives the same price for payer and receiver options - so this is in line with what we defined as the ATM Forward in equation 68, i.e. it is the default adjusted forward index value. On the CDSO screen it is given as 107.1444 which corresponds to a forward default adjusted value of -$7,144,443.\(^{64}\)

We calculate this value a few different ways: The first is to use the prices given in table 10 to build a pseudo credit credit for the index.\(^{65}\) We may then use equation 69 to compute the ATM forward price. The second method is to compute the intrinsic value from the credit curves of the constituent entries of the index using equation 68. The third method is to first adjust the constituent credit curves to match the index prices (as described in section 5.5.2), then compute the intrinsic value. We consider the last of these to be the correct way to compute the ATM forward. Table 11 shows Bloomberg’s value together with the three values we calculate. The unadjusted intrinsic value shows the largest difference (0.316%), while the homogenous pool and the adjusted intrinsic value show very close agreement with each other (a difference of less than $600 on a $100MM notional), and close agreement with Bloomberg (a difference of less than $6000 or 0.08%).

<table>
<thead>
<tr>
<th>ATM Forward</th>
<th>Error</th>
</tr>
</thead>
<tbody>
<tr>
<td>Bloomberg</td>
<td>-$7,144,443.00</td>
</tr>
<tr>
<td>Curve Built from Index Prices</td>
<td>-$7,149,840.40</td>
</tr>
<tr>
<td>Intrinsic Value</td>
<td>-$7,121,885.71</td>
</tr>
<tr>
<td>Adjusted Intrinsic Value</td>
<td>-$7,150,407.60</td>
</tr>
</tbody>
</table>

Table 11: Value of the ATM Forward (derived) from CDSO together with the calculated values from three methods described in the main text. The error shows percentage difference from the Bloomberg value.

There is nothing in either the documentation [FNS11] or on the CDSO screens to indicate that Bloomberg considers the underlying credit curves in the computation of index options. Most likely they assume a homogenous pool and build a single credit curve from the prices of different terms of the index. That our version of this calculation does not exactly match is most likely down to the fact they use their own model to build the index curve in this case.

9.3 Option Prices

We price a set of these index options using our implementation of the Pedersen model presented in section 6.7.1. Firstly we use the ATM forward value \(F^P(t, T_e, T_m)\) shown on the CDSO screen - we call this calculation 1. Secondly we use our own calculated value of \(F^P(t, T_e, T_m)\) from building an index curve - we call this calculation 2. We could have used the adjusted intrinsic value, but as we showed above, the difference in this case is tiny. Table 12 shows the results of these calculations.

Even when we use the same forward as Bloomberg (so there is no issue with curve construction or whether the constituent curves are considered) our option prices differ from Bloomberg’s. The

\(^{64}\)We have backed out the extra accuracy from the put-call relationship.

\(^{65}\)We use our implementation of the ISDA model that allows this. See [Whi13] for details.
Table 12: Prices of options on CDX.NA.HY index. See main text for details.

<table>
<thead>
<tr>
<th>Strike</th>
<th>Bloomberg Payer</th>
<th>Bloomberg Receiver</th>
<th>Calculation 1 Payer</th>
<th>Calculation 1 Receiver</th>
<th>Calculation 2 Payer</th>
<th>Calculation 2 Receiver</th>
</tr>
</thead>
<tbody>
<tr>
<td>103.000</td>
<td>1050.57</td>
<td>4144889.03</td>
<td>1166.48</td>
<td>414837.71</td>
<td>1150.74</td>
<td>4150217.96</td>
</tr>
<tr>
<td>104.000</td>
<td>7281.74</td>
<td>3151266.17</td>
<td>7845.01</td>
<td>315702.55</td>
<td>7755.71</td>
<td>3157009.24</td>
</tr>
<tr>
<td>105.000</td>
<td>38630.22</td>
<td>2182760.61</td>
<td>40556.79</td>
<td>2184600.64</td>
<td>40180.03</td>
<td>2189619.88</td>
</tr>
<tr>
<td>106.000</td>
<td>154451.94</td>
<td>1298728.30</td>
<td>158851.69</td>
<td>1303081.86</td>
<td>157696.67</td>
<td>1307322.84</td>
</tr>
<tr>
<td>107.144</td>
<td>529900.91</td>
<td>529900.91</td>
<td>536240.00</td>
<td>536240.00</td>
<td>533468.84</td>
<td>538864.84</td>
</tr>
<tr>
<td>108.000</td>
<td>1053705.97</td>
<td>198274.25</td>
<td>1058884.29</td>
<td>203487.10</td>
<td>1054818.41</td>
<td>204817.21</td>
</tr>
<tr>
<td>109.000</td>
<td>1894442.50</td>
<td>39156.75</td>
<td>1896630.82</td>
<td>41419.94</td>
<td>1891603.29</td>
<td>41788.41</td>
</tr>
<tr>
<td>110.000</td>
<td>2859098.57</td>
<td>3958.78</td>
<td>2859437.34</td>
<td>4412.77</td>
<td>2854092.97</td>
<td>4464.40</td>
</tr>
<tr>
<td>111.000</td>
<td>3855162.16</td>
<td>168.34</td>
<td>3855042.43</td>
<td>204.19</td>
<td>3849649.48</td>
<td>207.23</td>
</tr>
</tbody>
</table>

variation due to which method is used to convert between the (pseudo) spread and the forward price (i.e. the ISDA model or an approximation involving the credit triangle) is too small to explain this difference - when we change from the approximation discussed in section 7 to using the full ISDA model, the price of a receiver with an exercise price of 111 changes from 204.19 to 204.64, which is one-eighth of the discrepancy we have with Bloomberg's number.

Furthermore, the option prices we compute with our own value of the ATM forward (calculation 2) differ far less from calculation 1 than from the Bloomberg numbers - it is a theme of this section, that while we are confident in our implementation, as we have set out in this paper, we cannot hope to exactly reproduce a black box such as Bloomberg.

An alternative way of looking at the discrepancy is shown in figure 10. There we show the implied volatility corresponding to the option prices given on CDSO - that is, the volatility we must put in our model to obtain Bloomberg prices (recall the volatility is set to 30%). Again a 40bps difference in implied volatility is large, given that the underlying models should be the same up to some implementation details.

For completeness we should mention the role of the discount factor, $P(t,t_\text{es})$. We compute this value as 0.999814 (from our bootstrapped yield curve), however what we back-out from Bloomberg’s option prices is 0.99985 - in terms of zero rates 66 this is 0.174% versus 0.137%. Since the 1M and 2M deposit rates are 0.15% and 0.2% respectively (see table 10), the second number seems a little low. What this means is that the put-call relationship (eqn. 77) holds for all the sets of prices provided the ‘correct’ discount factor is used. The difference in discount factors does not account for the difference seen in the option prices.

9.3.1 Comparison of Greeks

In our Greeks section we used options on the 5Y CDX.NA.HY.21-V1 as examples. Here we compare our calculations with the values obtained from CDSO. Table 13 shows the full set of Greeks available on CDSO together with our calculations (most of which we show in section 8). For CS01 we use the flat spread method; delta and Gamma use CS01 ratio; Vega is for a 1% bump in (log-normal) volatility; Theta is the (normal) with defaults Theta over one day; and IR DV01 was computed with a flat index spread.

66These are given as continuously compounded.
We have already seen that our options prices do not agree, so we would not expect an exact agreement on the Greeks. With that caveat, the numbers are close. The notable exception is the Theta of payer options: The values from CDSO for deeply in and out the money options (where there is little optionality) is around -$8000 rather than zero - in fact the entire Theta curve looks like it is pulled down by $8000. This is not just for this index; we find the same thing for options on iTraxx Xover (with the same magnitude) and options on CDX.NA.IG. The options we have considered have more than a month to expiry, yet on the CDSO screen it is easy to have a Theta (the one day change in the option price) that is greater than the current option price - this makes no sense to us.

9.4 Index Options - iTraxx Xover

Our second index option example is on 5Y iTraxx Xover.\footnote{iTraxx Europe Crossover Series 20 version 1 5Y.} Like most indices, this is spread based, so the strikes are given as spreads which must be converted to actual exercise prices. The notional is €100MM, the trade date is the 6-Feb-2014, the option expiry is 19-Feb-2014, and the maturity of the underlying index is 20-Dec-2018. The index has a coupon of 500 bps and a recovery rate of 40%. The current spread is 318.25 bps which corresponds to PUF of -7.768. The market data used to build the yield curve (ISDA Standard EUR) and the spreads for other terms of iTraxx Xover are given in table 14.

9.4.1 Exercise Price

There are zero accrued days on 19-Mar-2014, so the full exercise amount is just the clean price of the index at $T_e$ (with index factor 1, valued at $t_{ex}$) for a quoted spread of $K$. This can be calcu-
Table 13: Comparison of Greeks from options on CDX.NA.HY

<p>| Strike Price | Bloomberg | | | | | | OpenGamma | | | | | |
|--------------|-----------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|
|              | CS01      | Delta           | Gamma           | Vega            | Theta           | IRDV01          | CS01            | Delta           | Gamma           | Vega            | Theta           | IRDV01          |
| 103          | 132.90    | 0.28%           | 0.2800%         | 419.00          | -8413.00        | 2.00            | 138.95          | 0.31%           | 0.4941%         | 450.84          | -166.32         | 1.53            |
| 104          | 763.54    | 1.64%           | 1.9262%         | 1885.95         | -8994.79        | 14.99           | 781.47          | 1.72%           | 2.0107%         | 1980.10         | -787.08         | 12.02           |
| 105          | 3264.81   | 7.03%           | 5.6222%         | 6026.67         | -10738.22       | 77.63           | 3269.04         | 7.18%           | 5.7525%         | 6200.47         | -2610.19        | 65.56           |
| 106          | 10086.19  | 21.72%          | 10.8710%        | 12962.36        | -13776.41       | 283.33          | 9939.58         | 21.83%          | 10.9853%        | 13160.42        | -5754.79        | 247.55          |
| ATM          | 24225.91  | 52.17%          | 12.8024%        | 17633.09        | -15816.23       | 797.40          | 23642.68        | 51.94%          | 12.9199%        | 17841.31        | -7928.54        | 716.91          |
| 108          | 35412.11  | 76.26%          | 8.9212%         | 13946.80        | -14069.72       | 1275.26         | 34507.50        | 75.80%          | 9.1124%         | 14231.49        | -6259.34        | 1166.39         |
| 109          | 43518.12  | 93.72%          | 3.1267%         | 5843.59         | -10424.58       | 1673.27         | 42497.16        | 93.35%          | 3.3217%         | 6120.34         | -2580.19        | 1553.41         |
| 110          | 46073.15  | 99.22%          | 0.4589%         | 1143.52         | -8454.46        | 1814.60         | 45100.02        | 99.07%          | 0.5631%         | 1258.43         | -480.55         | 1696.57         |
| 111          | 46438.91  | 100.01%         | 0.0207%         | 89.13           | -8086.11        | 1831.47         | 45494.39        | 99.94%          | 0.0379%         | 106.44          | -20.09          | 1713.34         |
| 103          | -46327.39 | -99.77%         | 0.5136%         | 418.59          | -395.36         | -1895.46        | -45379.86       | -99.69%         | 0.4941%         | 396.57          | -148.17         | -1800.47        |
| 104          | -45966.03 | -98.41%         | 1.9784%         | 1885.95         | -932.64         | -1874.50        | -44737.34       | -98.27%         | 2.0107%         | 1834.10         | -773.30         | -1779.14        |
| 105          | -43194.76 | -93.02%         | 5.6744%         | 6026.67         | -2631.68        | -1803.81        | -42249.77       | -92.81%         | 5.7525%         | 5974.56         | -2600.78        | -1714.78        |
| 106          | -36373.38 | -78.33%         | 10.9233%        | 12962.36        | -5625.48        | -1590.06        | -35579.22       | -78.16%         | 10.9854%        | 13015.79        | -5749.76        | -1521.96        |
| ATM          | -22233.66 | -47.88%         | 12.8546%        | 17633.09        | -7614.50        | -1066.77        | -21876.12       | -48.06%         | 12.9200%        | 17842.76        | -7928.51        | -1040.21        |
| 108          | -11047.46 | -23.79%         | 8.9735%         | 13946.80        | -5830.01        | -582.02         | -11011.31       | -24.19%         | 9.1124%         | 14136.90        | -6236.03        | -581.46         |
| 109          | -2941.45  | -6.33%          | 3.1790%         | 5843.59         | -2140.48        | -175.96         | -3921.64        | -6.64%          | 3.3218%         | 5915.92         | -2588.28        | -183.61         |
| 110          | -386.42   | -0.83%          | 0.5112%         | 1143.52         | -125.97         | 0.27            | -418.79         | -0.92%          | 0.5632%         | 1154.03         | -493.02         | -29.63          |
| 111          | -20.66    | -0.04%          | 0.0316%         | 89.13           | 286.78          | -1.65           | -24.42         | -0.05%          | 0.0380%         | 89.70           | -36.93          | -2.03           |</p>
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<td>5Y</td>
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<td>7Y</td>
<td>377.9800</td>
</tr>
<tr>
<td>10Y</td>
<td>401.3900</td>
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CDS Conventions

- Frequency: Quarterly IMM
- Day Count: ACT/360
- Recovery Rate: 40%

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<tr>
<td>Floating Interval</td>
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<tr>
<td>Bad Day Conv</td>
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</table>

Table 14: Market data used to build the Standard ISDA EUR yield curve (left), and Spreads on iTraxx Europe Crossover Series 20 for 6-Feb-2014 (right).
lated at $T_e$ using the ISDA model$^{68}$ with the yield curve calibrated at that time. The dependance on the yield curve is weak$^{69}$, so as usual, we calculate $G(K)$ at $t$ assuming deterministic interest rates.

For indices quoted on spread, the CDSO screen does not give the exercise price, so it must be backed out from the option prices. From equation 72 we have that

$$V_{\text{payer}}(t, T_e; T_m) - V_{\text{receiver}}(t, T_e; T_m) = P(t, t_{es})(F_D(t, T_e; T_m) - G(K))$$

(102)

The first term in brackets is the ATM forward value, while the second is the exercise price. Since $G(K) = 0$ if $K = C$ (the index coupon), we may extract the first term from the payer and receiver prices at a strike equal to the index coupon. We find this value to be -€7,195,599, which as expected, is close to the current (clean) value of the index. As the first term does not depend on $K$, we can then easily extract $G(K)$ for a range of $K$.

In Figure 11 we show the exercise price computed from the ISDA model along with the values we have backed out from options prices given on the CDSO screen. Given that all parties must agree on the exercise price, that there are significant differences$^{70}$ between the values we compute from the ISDA model and those we extract from the option prices, is a concern.

Figure 11: The exercise price for a index option for a range of strikes (given as basis point spreads). The crosses use the ISDA model to convert a spread to an upfront price, while the circles are the values extracted from Bloomberg option prices (see main text). The difference is shown as a line using the righthand scale.

$^{68}$ISDA Standard Upfront Model.

$^{69}$The IR DV01 is much smaller than the Credit DV01.

$^{70}$By construction the two exercise prices must agree for a strike of 500 bps (and have zero value).
9.4.2 ATM Forward Value

As in the previous example, the index is considered homogeneous and the index curve built from the quoted spreads at 3Y, 5Y, 7Y and 10Y. To do this properly, one should first convert the spreads to PUF and then bootstrap a credit curve that recovers all the upfront amounts.

Once this curve is built we may compute the default adjusted index value from equation 69. We obtain a value of -€7,201,983, while, as we saw above, the value derived from CDSO is -€7,195,599, so there is a discrepancy of €6,385 (about 0.09%), which again we cannot reconcile without further details of how the value is obtained, which depends, among other things, on exactly how the credit curve is build.

9.4.3 ATM Forward Spread

The ATM forward given on CDSO is 331.79bps. However this is not the value of the strike that makes the price of the payer and receiver equal - that is a strike of 326.61bps. The value we calculate using equation 78 (which we call the ATM forward spread, $K_{ATM}$) is 327.39bps - a discrepancy of about 0.8bps from the value of 326.61bps we extracted from CSDO. A calculation of the default-adjusted forward spread, $S^D$, using equation 81 gives a value of 331.65bps - this is within 0.15bps of the Bloomberg value for the ATM forward.

It seems reasonable to assume that Bloomberg uses (a variant of) equation 81 to calculate what it calls ATM Forward [spread]. However this number is not the ATM forward spread, rather it is the defaulted-adjusted forward spread.

9.4.4 Option prices

We price these index options using our implementation of the Pedersen model presented in section 6.7.1. Firstly we use the value of $F^D(t, T_e, T_m)$ and the exercise prices $G(K)$ derived from the option prices on the CDSO screen - we call this calculation 1. Secondly we use our own calculated values of $F^P(t, T_e, T_m)$ and $G(K)$ - we call this calculation 2. Table 15 shows the results of these calculations. The first set (calculation 1), unsurprisingly shows close agreement with Bloomberg, while the second set (calculation 2) show a larger discrepancy. In this second case we have used a different value of the ATM forward value (which is used to calibrate the model parameter $X_0$) and different values for the exercise prices, so it would be surprising to obtain the same option values.

The pertinent question is which values are correct. The calculation of default adjusted index value depends on whether one builds a single index curve or adjusts all the consistent curves, and in either case exactly how this is done. The value is therefore somewhat subjective. The exercise price on the other hand is an amount that is actually paid if the option is exercised, and thus there must be an agreed procedure to convert between a strike given as a spread and this amount. Since for spot indices this procedure is the ISDA standard model, it would seem reasonable to use the same procedure to calculate the excise amount. However Bloomberg does something different which results in different exercise prices.

\footnote{We obtained this value by manually root finding.}
Table 15: Prices of options on ITX Xover. See main text for details.

A Mathematical Preliminaries

A.1 Optional Pricing Theory

We state here without proof a couple of useful results from Martingale pricing theory that we use in the main text. For details see one of the many books that cover this material in a financial context (for example [HK04, Shr04, Reb02, RMW09]).

Let \( X(u) \) be the value of some asset at time \( u \), that depends only on the value of some (finance) observables at time \( u \). Let \( N(u) \) be the value of some other asset at time \( u \), such that \( N(u) > 0 \ \forall \ u < T \), where \( T \) is some horizon time. We may measure the value of \( X \) in terms of \( N \), i.e.

\[
\tilde{X}(u) = \frac{X(u)}{N(u)}
\]

The equivalent Martingale theory states that there exists a probability measure \( P^N \) such that \( \tilde{X}(u) \) is a Martingale. That is

\[
\mathbb{E}_t^N [ \tilde{X}(u) ] = \mathbb{E}_t^N \left[ \frac{X(u)}{N(u)} \right] = \tilde{X}(t) = \frac{X(t)}{N(t)}
\]

(103)

for \( u \geq t \). \( \mathbb{E}_t^N [\cdot] \) is shorthand for \( \mathbb{E}_t^N [\cdot | F_t] \), which is the expectation under the \( P^N \) measure conditional on the filtration \( F_t \), which loosely means that the expectation only depends on information available at \( t \). \( N(\cdot) \) is known as the numeraire associated with the measure \( P^N \).

The value today (time \( t \)) of a derivative with payoff at \( T \) of \( X(T) \) is given by

\[
V(t) = \mathbb{E}_t \left[ e^{-\int_t^T r_s ds} X(T) \right] = \beta_t(t) \mathbb{E}_t \left[ \frac{X(T)}{\beta_t(T)} \right]
\]

(104)
where \( r_s \) is the risk-free (instantaneous) rate at \( s \) and \( \beta_t(s) = e^{\int_s^t r_s \, ds} \) is the value of a money market account started at \( t \) (so \( \beta_t(t) = 1 \)), i.e. it is the numeraire associated with the risk-neutral measure.\(^2\) We can use the result of equation 103 to write the option price as

\[
V(t) = N(t) \mathbb{E}_t^N \left[ \frac{X(T)}{N(T)} \right]
\]

(105)

This is useful in the case that the payoff can be written in the form \( X(T) = N(T)Y(T) \). Then

\[
V(t) = N(t) \mathbb{E}_t^N [Y(T)]
\]

(106)

where \( Y(\cdot) \) is a Martingale under \( \mathbb{P}^N \). Then any model form the dynamics of \( Y(\cdot) \) that is a Martingale under \( \mathbb{P}^N \), will produce arbitrage free prices.

### A.2 Curves

For a given \( t \), \( P(t, T) \) and \( Q(t, T) \) are functions (or curves) of the independent variable \( T \), and are known as the discount (or yield) curve and the credit curve. They form the bases of CDS pricing \([\text{Whi13, O'k08}]\).

#### A.2.1 Discount Curve

We assume that at an observation time, \( t \), the price of a risk-free\(^7\) zero-coupon bond that pays 1 at time \( T \geq t \) is \( P(t, T) \). This is also known as the (risk-free) discount factor, since we can discount a known cash flow of \( c \) at time \( T \) to today (time \( t \)) by selling \( c \) zero-coupon bonds for \( cP(t, T) \). Clearly \( P(T, T) = 1 \).

Define \( P(t, S, T) \), \( T > S \geq t \), as the forward discount factor observed at \( t \) for the period \( S \) to \( T \). That is, I agree today (\( t \)) to pay \( P(t, S, T) \) at time \( S \) to receive 1 at time \( T \). By a standard no arbitrage argument, we have

\[
P(t, S, T) = \frac{P(t, T)}{P(t, S)}
\]

(107)

and \( P(S, S, T) = P(S, T) \). Since interest rates evolve over time, generally \( P(t, S, T) \neq P(S, T) \). However in the \( S \) forward measure which uses \( P(\cdot, S) \) as the numeraire, we have

\[
\mathbb{E}_t^S [P(u, S, T)] = P(t, S, T) \quad \forall u \geq t
\]

(108)

since the \( P(u, S, T) \) is the ratio of an asset to the numeraire. So \( P(u, S, T) \) is a Martingale in the \( S \) forward measure.

#### A.2.2 Credit Curve

Let \( \tau \) be the default time of a particular obligor. If it has not defaulted by \( t \) (i.e. \( \tau > t \)), then the probability that it has not defaulted by \( T \geq t \) is

\[
Q(t, T) = \mathbb{E}_t[I_{\tau > T} \mid \tau > t] = P(\tau > T \mid \tau > t)
\]

(109)

in the risk-neutral measure. If default has occurred (i.e. \( \tau \leq t \)), then \( Q(t, T) = 0 \).

\(^2\)Often this is written as \( \mathbb{E}_t^N [\cdot] \) - we suppress the superscript when dealing with the risk-neutral measure.

\(^7\)There is no chance that the counterpart will fail the pay.
Define $Q(t, S, T)$ as the probability observed at $t$ that the obligor will not have defaulted by $T$ conditional on it not defaulting by $S$. It is a standard conditional probability result that

$$Q(t, S, T) = \frac{Q(t, T)}{Q(t, S)}.$$  

(110)

As the credit quality of the obligor will evolve over time, have $Q(t, S, T) \neq Q(S, T)$.

### A.3 Survival Measure Pricing

Consider a derivative that pays an amount $Z(T)$ (which is fixed at $T$) at some time $t$ if default occurs after $T$, but before $t$, the payment is still made. The payoff at $T$ is $P(T; t \geq T)$ and the present value is given by

$$Z(t; T) = \mathbb{E}_t \left[ P(T; t \geq T) Z(T) 1_{t > T} \right]$$

Assume that we can calculate the value $Z(t; T)$ by some means. Now consider a second derivative that pays an amount $X(T)$ at $t$ conditional on the same reference obligor not defaulting by $T$. Its present value is give by

$$X(t; T) = \mathbb{E}_t \left[ \frac{P(T; t \geq T)}{\beta_t(T)} X(T) 1_{t > T} \right]$$

We would like to price this by using the first asset, $Z(t; T)$, as a numeraire, but since its value becomes zero in the event of a default, it fails the technical condition that $N(u) > 0 \forall u < T^*$. However, we are saved by the fact that the derivative value, $X(t; T)$ also becomes zero in the event of a default, so there is no technical restriction. This means we are free to change to the $Z$ survival measure [Sch04] with $Z(t; T)$ as the numeraire and write

$$X(t; T) = Z(t, T) \mathbb{E}_T^Z \left[ \frac{P(T; t \geq T) X(T) 1_{t > T}}{P(T; t \geq T) Z(T) 1_{t > T}} \right] = Z(t, T) \mathbb{E}_T^Z \left[ \frac{X(T)}{Z(T)} \right]$$

(111)

The survival indicator has disappeared from the expectation. This is particularly useful if we can write $X(T) = Z(T) Y(T)$, which gives

$$X(t; T) = Z(t; T) \mathbb{E}_T^Z \left[ Y(T) \right]$$

so the task is to compute the expectation of $Y(T)$ under $Z$.

#### A.3.1 Option Payoff

Imagine an option payoff at $T$ of the form

$$p(T; t \geq T) (\chi[X(T) - K Z(T)])^+ = p(T; t \geq T) 1_{t > T} (\chi Z(T) [Y(T) - K])^+$$

We may compute its present value thus

$$V_{\text{option}}(t; T) = Z(t; T) \mathbb{E}_T^Z \left[ (\chi [Y(T) - K])^+ \right]$$

Since $Y$ is a Martingale in the $Z$ survival measure, any model that preserves this property will produce arbitrage free prices. In particular the Black model can be used in this case.
A.3.2 The Risky Bond Measure

A risky (zero coupon) bond issued by some obligor will pay 1 at time $t_s$ if the obligor has not defaulted by $T$.\(^{74}\) Its present value is

$$B(t, T) = E_t \left[ e^{-\int_t^{t_s} r_s ds} I_{t_s > T} \right]$$

(112)

If we assume that credit events are independent of interest rates,\(^{75}\) we may write

$$B(t, T) = E_t \left[ e^{-\int_t^{t_s} r_s ds} E_t \left[ I_{t_s > T} \right] = P(t, t_s)Q(t, T) \right]$$

(113)

However, it is not necessary to assume the independence of credit events and interest rates. We may change to the $T$ forward measure (with $P(t, T)$ as the numeraire), which gives

$$B(t, T) = P(t, T)E^T_t \left[ P(T, t_s) I_{t_s > T} \right]$$

(114)

If $T = t_s$ (i.e. there is no payment delay) and we redefine the survival probability as

$$Q(t, T) = E_T^T \left[ I_{t_s > T} \right]$$

(115)

then

$$B(t, T) = P(t, T)Q(t, T)$$

with no independence assumption - all we have done is define $Q(t, T)$ as the survival probability in the $T$-forward measure, rather than the risk-neutral measure.\(^{76}\)

In practice when a payment delay exists, it will be only a few days (i.e. three working days from the exercise of an option to settlement of payments), so we may write

$$E_T^T \left[ P(T, t_s) I_{t_s > T} \right] \approx E_T^T \left[ P(T, t_s) \right] E_T^T \left[ I_{t_s > T} \right] = \frac{P(t, t_s)}{P(t, T)} \left[ I_{t_s > T} \right]$$

which gives

$$B(t, T) = P(t, t_s)Q(t, T)$$

The function $B(t, \cdot)$ is known as the risky discount curve (for a particular obligor and currency).

This risky bond is equivalent to setting $Z(T) = 1$ above. This means a derivative with payoff $X(T)$ (at $t_s$) conditional on the bond not defaulting, has a present value given by

$$V(t) = B(t, T)E_T^T \left[ \frac{P(T, t_s)}{B(T, T)} X(T) I_{t_s > T} \right] = B(t, T)E_T^T \left[ X(T) \right]$$

(116)

since $B(T, T) = P(T, t_s)I_{t_s > T}$. To price the derivative we would need to compute the expectation of $X(T)$ under the risky bond measure.

A.4 The Annuity Measure

Consider a derivative price that can be written as

$$V(t, T) = E_t \left[ e^{-\int_t^T r_s ds} A(T, S)X(T) I_{t_s > T} \right]$$

(117)

$$= E_t \left[ \frac{P(T, t_s)}{\beta_T(t)} A(T, S)X(T) I_{t_s > T} \right]$$

---

\(^{74}\)Assume the money is held in escrow between $T$ and $t_s$, so it is still paid if there is a default in this period.

\(^{75}\)Since an increase in debt cost is a contributing factor in bankruptcy, they are not really independent.

\(^{76}\)It follows that if rates and credit events are independent, then $E_T^T \left[ I_{t_s > T} \right] = E_t \left[ I_{t_s > T} \right]$. 52
where \(A(T, S)\) is the annuity at \(T\) (for some CDS maturity \(S > T\)), with cash settlement at \(t_{cs}\). Since the annuity is an asset (the premium leg of a unit coupon), we may set \(Z(T) = A(T, S)\) and change to the *annuity measure* \(A\). This gives the derivative price as

\[
V(t, T) = A(t, T, S) \mathbb{E}_t^A \left[ \frac{P(T,t_{cs})A(T, S)1_{t>T}}{P(T,t_{cs})A(T, S)1_{t>T}} X(T) \right]
\]

\[= A(t, T, S) \mathbb{E}_t^A [X(T)] \tag{118}\]

where

\[
A(t, T, S) = \mathbb{E}_t \left[ \frac{P(T,t_{cs})}{\beta(t)} A(T, S)1_{t>T} \right]
\]

Any quantity defined as

\[
Y(t, T) = \frac{X(t, T)}{A(t, T, S)}
\]

will be a Martingale in the annuity measure, i.e. \(\mathbb{E}_t [Y(u, T)] = Y(t, T) \forall u \geq t\). This includes the forward spread.

B **List of Terms**

B.1 **General Terms**

- \(t\) today or the trade date.
- \(t_{cs}\) Cash settlement time for a spot CDS/index. For standard CDS and CDS indices this is three working days after the trade date.
- \(T_e\) The exercise date of a CDS/index option or the forward start date of a forward starting CDS/index.
- \(t_{es}\) The exercise settlement date. This would normally correspond to the cash settlement date of a spot CDS with trade date \(T_e\) (i.e. \(t_{es}\) is three workings days after \(T_e\)).
- \(T_m\) The maturity (protection end) of a CDS.
- \(s_i\) The start of the \(i\)th accrual period.
- \(e_i\) The end of the \(i\)th accrual period.
- \(t_i\) The \(i\)th payment date.
- \(\Delta_i\) The year fraction of the \(i\)th accrual period.
- \(\Delta\) The year fraction in the current accrual period, i.e. the year fraction between \(s_1\) and \(t+1\) (the spin-in date). Multiplied by the day count convention base (almost always 360) this gives the accrued days. Multiplied by the coupon and the notional, it gives the (magnitude) of the accrued. Also used for the year fraction at \(T_e\) for a forward starting CDS (i.e. one starting at \(T_e\)).
- \(C\) The coupon of a CDS.
• PUF Points Up-Front. The clean price of a CDS or index quoted as a percentage of notional.
• N The (initial) notional of a CDS or index.
• \( r_s \) Risk-free instantaneous interest rate at time \( s \).
• \( P(t, s) \) The price at \( t \) of a risk free zero coupon bond that pays 1 at \( s \geq t \). \( P(t, \cdot) \) is the discount curve observed at \( t \).
• \( Q(t, s) \) The survival probability - the probability that a reference entity that has not defaulted by \( t \) will not have defaulted (i.e. survived) by \( s \geq t \). \( Q(t, \cdot) \) is the survival or credit curve observed at \( t \).
• \( h(t, s) \) The forward instantaneous hazard rate. This is related to the survival probability by \( Q(t, s) = \exp\left(- \int_t^s h(t, s') \, ds' \right) \)
• \( \dot{Q}(t, s) \equiv \frac{dQ(t, s)}{ds} = -h(t, s)Q(t, s) \). This is the (negative of) the probability density for default at \( s \) given no default by \( t \leq s \).
• \( \Lambda(t, s) \) The zero hazard rate. This is related to the survival probability by \( Q(t, s) = \exp(-\Lambda(t, s)(s - t)) \)
• \( \mathbb{E}_t[\cdot] \) Risk neutral expectation given the information available up to time \( t \).
• \( \mathbb{E}_X^X[\cdot] \) Expectation under some equivalent measure \( X \) given the information available up to time \( t \).
• \( \beta_i(s) = \exp\left(\int_t^s r_u \, du \right) \) is the money market numeraire; It is the amount of money in a ‘risk-free’ bank account at time \( s \geq t \) if one unit of currency is invested at \( t \).
• \( \Delta X \) A small change in the quantity \( X \). This should not be confused with the use of \( \Delta \) for a year fraction.

B.2 Single Name CDS Terms
• \( \tau \) The default time. The random time when a reference entity defaults.
• \( \tau_{t \geq T} \) Default indicator. Takes the value 1 if the subscript is true, i.e. the default (\( \tau \)) is after \( t \), and 0 if it is false.
• \( R(\tau) \) Realised recovery rate for a default at \( \tau \).
• \( \tilde{R}(T) \) Expected recovery rate for protection to \( T \).
• \( R \) Expected recovery rate when there is no term structure of recovery (i.e. no dependence on length of protection).
• \( \eta_i \) Ratio of the year fraction for the \( i^{th} \) accrual period measured with the accrual day count convention to the same period measured with the curve day count convention. In almost all cases this is just 365/360.
• \( V_{perm}(t, T) \) Value of the premium leg of a CDS with protection end at \( T \) observed at \( t \). By convention the valuation is rolled forward to \( t_{ca} \geq t \).
• $V_{prot}(t, T)$ As above but for the protection leg.

• $A(t, T)$ The (risky) annuity, duration or RPV01 of a CDS with protection end at $T$. This is related to the premium leg by $V_{perm}(t, T) = C \times A(t, T)$.

• $S(t, T)$ The par spread of a CDS with maturity of $T$ - the value of the coupon that would make the fair value of the CDS zero. A set of these (corresponding to different maturities of CDS on the same name) can be used to construct a full credit curve; conversely, to computed this requires a full credit curve. By construction the following relationship holds: $V_{prot}(t, T) = S(t, T)A(t, T)$.

• A Constant hazard rate. A (often hidden) parameter used to convert between PUF and a quoted spread for a CDS or index.

• $S(t, T)$ The quoted or flat spread of a CDS maturity $T$. If a CDS is priced assuming a constant hazard rate, then all quantities we calculate from that ‘flat’ credit curve, we prefix with flat and use an over-bar in the maths. The flat spread is the spread that is calculated from a flat credit curve.

• $V_{prot}(t, T)$ or $\tilde{V}_{prot}(t, T|\tilde{S})$ The value of the flat protection leg of a CDS. The value of the protection leg calculated using a flat (i.e. constant hazard rate) credit curve. See above.

• $A(t, T)$ or $\tilde{A}(t, T|\tilde{S})$ The flat annuity. The value of the annuity calculated using a flat (i.e. constant hazard rate) credit curve (see above). In the second case we mean the annuity corresponding to the flat spread $\tilde{S}$.

• $V(t, T)$ Value of a CDS with protection end (maturity) at $T$ observed at $t$. By convention the valuation is rolled forward to $t_{cs} \geq t$.

• $V(t, T_e, T_m)$ Present value of a forward starting CDS. The CDS is entered at $T_e$ and protection ends at $T_m > T_e$.This value includes the possibility of default before $T_e$ (in which case the contract is worth zero - there is no frontend protection).

• $V_{prot}(t, T_e, T_m)$ Present value of the protection leg of a forward starting CDS. See above.

• $A(t, T_e, T_m)$ Present value of the annuity of a forward starting CDS. See above.

• $F(t, T_e, T_m)$ Forward value of CDS. The expected value of a CDS at $T_e$ conditional on the reference not having defaulted by $T_e$. The CDS maturity is $T_m$ and the expecting is calculated with all the information available at $t$.

• $S(t, T_e, T_m)$ The forward spread observed at $t$ for protection starting at $T_e > t$ and ending at $T_m > T_e$.

• $\tilde{S}_0$ Expected value of the flat spread $\tilde{S}(T_e, T_m)$ at $t$ under the risky bond measure.

• $K_{ATM}(t, T_e, T_m)$ The ATM forward flat spread defined in section 3.5.1.
B.3 CDS Index Terms

- $J$ The initial number of entries in an index.
- $J_D$ The number of defaults from an index. $J_D(t)$ the number of defaults at time $t$.
- $w_i$ The weight of the $i^{th}$ name in an index. For an equally weighted index this is just $1/J$.
- $f$ The index factor - the fraction of names remaining in an index. $f(t)$ The index factor at time $t$.
- $R_I$ Recovery rate of an index. This is purely used to convert between the index spread and an upfront amount.
- $V_I(t, T)$ (Clean) value of an index with maturity $T$.
- $V_{I,pro}t(t, T)$ Nominal value of the protection leg of an index using a constant hazard rate.
- $A_I(t, T)$ Flat annuity of an index, calculated from the (flat) spread, $S$.
- $\tilde{S}_I(t, T)$ Index flat spread. An alternative way of expressing the price of an index.
- $V_{I,pro}t(t, T)$ Value of the protection leg of an index, calculated using adjusted intrinsic credit curves.
- $A_I(t, T)$ Annuity of an index, calculated using adjusted intrinsic credit curves.
- $S_I(t, T)$ Index spread, calculated using adjusted intrinsic credit curves.
- $V_D^I(T_e, T_m)$ The clean value of an index plus the settlement value of any defaults from the index before $T_e$. This is known as the default-adjusted forward portfolio swap.
- $F_D^I(t, T_e, T_m)$ The ATM forward value or default-adjusted forward index value. Given by $F_D^I(t, T_e, T_m) \equiv \mathbb{E}^T_t \left[ V_D^I(T_e, T_m) \right]$.
- $S_I(t, T_e, T_m)$ The forward index spread - this does not include default settlement.
- $S_D^I(t, T_e, T_m)$ The default-adjusted forward index spread.
- $K_{ATM}$ The ATM forward spread. It is the strike (spread) that makes the exercise price equal to the The ATM forward value.

Additionally, any single-name term, $X$ when considered as the $j^{th}$ entry of an index is given the notation $X_j$.

B.4 Option Terms

- $V_{\text{option}}(t, T_e, T_m)$ The price of an option (payer or receiver) at time $t$ on a single-name CDS with or without a knockout feature. The expiry is at $T_e$ and the maturity of the underlying CDS is $T_m$.
- $V_{I,\text{option}}(t, T_e, T_m)$ The price of an option (payer or receiver) at time $t$ on a CDS index. The expiry is at $T_e$ and the maturity of the underlying index is $T_m$. 

• \( \chi \) Takes the value +1 for a payer option and −1 for a receiver option.
• \( \sigma(t) \) Time dependent, deterministic volatility. \( \hat{\sigma} \) or \( \sigma \) is root-mean-square (RMS) value.
• \( K \) Strike level expressed as a spread.
• \( G(K) \) The exercise price of an option. This is often a function of the strike (spread) \( K \).
• \( \text{Black}(F, K, T, \sigma, \chi) \) The Black formula (see eqn 40).
• \( \bar{X} \) A flat pseudo spread that absorbs defaults from the index. Used in the Pedersen model.
• \( \bar{X}_0 \) The calibrated mean of \( \bar{X} \).
• \( \tilde{K} \) The modified strike used the modified Black index option pricing formula (eqn. 87).

References


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